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# MOMENT PROPAGATION METHODS FOR STOCHASTIC SIMULATION OF COMPLEX BIOCHEMICAL SYSTEMS 

A Dissertation Presented<br>by<br>Vibha Mane<br>to<br>The Graduate School<br>in Partial Fulfillment of the<br>Requirements<br>for the Degree of<br>Doctor of Philosophy<br>in<br>Electrical Engineering<br>Stony Brook University

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# Abstract of the Dissertation <br> MOMENT PROPAGATION METHODS FOR STOCHASTIC SIMULATION OF COMPLEX BIOCHEMICAL SYSTEMS 

by

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We are interested in predicting the time dependent behavior of biochemical networks such as interaction between proteins. These networks are represented by a system of chemical reactions. In the forward problem, we have knowledge of all the reactions and the associated rate constants. We want to determine the joint probability density of the populations of all the molecular species at any time instant.

The given biochemical system is a discrete state, continuous time Markov process, and the time evolution of its probability density function is described by the chemical master equation (CME). We want to obtain the solution of the master equation in complex biochemical systems with a large number of species and reactions. Analytical solutions of the chemical master equation for first and second order reactions have only been obtained for select cases with a few species and reactions. Another approach to solving the problem is
to approximate the CME with the Fokker-Planck equations. But this would require solving partial differential equations with a large number of variables.

The present state of the art approaches for stochastic simulation of such systems are based on Monte Carlo methods. One such popular method is the Stochastic Simulation Algorithm (SSA) derived by Gillespie in 1976. Several authors have developed accelerated versions of the SSA such as the Next Reaction Method and the $\tau$-leap methods, in order to reduce the computation time of SSA. The Monte Carlo methods provide approximations of the complete distribution, but they require simulations of many realizations of the Markov process and many time steps. Hence their computation times are prohibitively long for very complex systems.

Methods for modeling the biochemical networks based on moment propagation is a relatively unexplored area. We propose a new method for propagating the first two moments of the joint probability distribution of the number of molecules. In many systems, the distribution can be approximated as Gaussians and therefore computing the first two moments is sufficient. Simulation results show that our method yields accurate results for first order and second order reactions. Compared with the Monte Carlo methods, our method yields significant savings in computation time. Compared with other moment propagation methods, the recursive expressions in our method can be implemented by specifying rate constants and stoichiometries, without having to derive or solve any differential equations. Whereas other moment propagation methods with similar accuracy have been demonstrated for a few species, we demonstrate our method for complex biochemical systems with hundreds of species.

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## Chapter 1

## Introduction

### 1.1 Introduction

Molecular and systems biologists aim to develop a better understanding of complex interactions in biological systems. These include interactions between DNA, RNA and proteins in a gene regulatory system, enzyme-substrate interactions in metabolic pathways and interactions between proteins in cell signaling pathways. We start the modeling of a biological process by representing it with a system of chemical reactions. We next apply a computational model to predict the behavior of the system. In the deterministic framework, we have a set of reaction rate equations with the state variable being the concentration of molecular species. The deterministic approach adequately models many macroscopic systems. The underlying assumption in this approach is that the molecular populations are very large and can be approximated as continuously varying quantities. However, in intracellular biochemical reactions molecular populations of some species are very small, and the deterministic approach fails to predict the system behavior quantitatively or qualitatively. It is well known that gene expression is a random phenomenon due to fluctuations in transcription, translation, and degradation of proteins and messenger RNA. This has been experimentally observed by Novick et al. [56], and has been described by Arkin and many others [2], [52], [53], [60]. Stochasticity in gene expression can give rise to phenotypically distinct subgroups with genetically identical cells [42].

In the stochastic framework, the state variable is a random vector representing the number of molecules of each species. Given the initial state of the system, the reaction
model and the associated rate constants, we want to obtain the joint probability mass function (PMF) of the state of the system at any later time. The time evolution of this joint PMF is described by the chemical master equation (CME), a differential-difference equation [25], [76]. In the analytical approach, the probability generating function is employed to transform the master equation into a partial differential equation (PDE) [54], [55]. Another approach is to utilize the Laplace transform to transform the master equation into a system of linear equations [46]. The analytical solution of the master equation has been obtained for select cases, with very few species and this will be described in Chapter 2. In the continuous stochastic formulation, the evolution of the joint probability density function (PDF) of the state of the system is described by the stochastic differential equation (SDE) such as the Fokker-Planck equation or the Langevin equation [28], [76].

The stochastic simulation algorithm (SSA), which is a Monte Carlo-based method was first introduced by Doob in 1945 [12], and derived by Gillespie in 1976 for the biochemical reaction system [22]. In this method, we generate many realizations of the stochastic process using random sampling, and average over these realizations to obtain the PMF of the state of the system. However, the SSA simulates only one reaction per time step, and hence it is not scalable with the population of molecular species, number of species or the number of reactions. Several authors have developed accelerated versions of the SSA such as the $\tau$-leap methods [9], [27], [58], [75]. The Monte Carlo methods will be discussed in Chapter 3.

Quite often, we are interested in the lower order moments of the PMF, such as the means and the covariance. In the moment propagation method, we trace a few lower order moments of the joint PMF of the state of the system. In 1995, Gillespie used the propagator function of a Markov process to derive moment evaluation equations [24]. In 2005, Hespanha et al. [38] and Singh et al. [67] presented a Stochastic Hybrid Systems (SHS) method to compute the first few moments of the state vector. In 2007, Gómez-Uribe et al. proposed a Mass Fluctuation Kinetics (MFK) method and obtained differential equations to estimate the means, variances and the covariances [32], [33]. These methods are evaluated in Chapter 6.

In the proposed method, a moment propagation method, our goal was to compute the evolution of the first two moments of the state vector for very large systems. We have utilized conditional expectation relations of random variables, Markovian property of the biochemical system, and expressions for simple reactions to build a general recipe for complex systems. We have demonstrated that the proposed method can be implemented
very easily by specifying rate constants and stoichiometries, without having to derive or solve any differential equations [50], [51]. The method is presented in Chapter 4, for first order reactions and Chapter 5 for second order reactions. Its comparison with the Monte Carlo methods and other moment propagation methods is provided in Chapter 6.

In the remainder of this chapter, we describe some biology examples where the stochastic simulation of molecular populations will be an additional tool to help comprehend such problems. Some of these examples are simulated in this research. We also discuss the deterministic method and introduce the stochastic approach.

### 1.2 Chemical Reactions

A chemical reaction such as

$$
\begin{equation*}
\nu_{11} S_{1}+\nu_{12} S_{2} \xrightarrow{k_{1}} \nu_{13} S_{3}+\nu_{14} S_{4} \tag{1.1}
\end{equation*}
$$

states that $\nu_{11}$ molecules of type $S_{1}$ react with $\nu_{12}$ molecules of type $S_{2}$ to form $\nu_{13}$ molecules of type $S_{3}$ and $\nu_{14}$ molecules of type $S_{4}$ [71]. The coefficients $\nu_{11}, \nu_{12}, \nu_{13}$, and $\nu_{14}$ are called stoichiometric coefficients of the reaction. In an elementary reaction, the above reaction occurs in a single step. In this report, a reaction written as in equation (1.1) will indicate an elementary reaction. The reaction rate tells us how fast a reaction takes place.

The rate equation relates the reaction rate to the concentration of reactants. For the chemical reaction in (1.1), the rate equation is

$$
\begin{equation*}
\mathrm{r}=k_{1}\left[S_{1}\right]^{\nu_{11}}\left[S_{2}\right]^{\nu_{12}} \tag{1.2}
\end{equation*}
$$

where $r$ is the reaction rate, $k_{1}$ is temperature dependent reaction coefficient, and [ $S_{1}$ ] and [ $S_{2}$ ] are the concentrations of the reactants $S_{1}$ and $S_{2}$, respectively. The units of concentration are mole litre ${ }^{-1}$ or mole $l^{-1}$. We also refer to the reaction coefficient $k_{1}$ as reaction constant or rate constant, although it may be time dependent in the computational models described henceforth.

For elementary reactions, we define reaction order as the sum of the powers of the concentrations in the reaction rate. The reactions are categorized as zeroth, first or second order reactions. Reactions of order three or more are very rare and may be modeled as a
sequence of second order reactions [30].
In the zeroth order reaction

$$
\varnothing \xrightarrow{k_{1}} S_{1}
$$

the reaction rate is independent of the concentration of the reactants and therefore for modeling purposes there are no reactants. Its rate equation is given by

$$
\mathrm{r}=k_{1}
$$

and the units of the reaction constant $k_{1}$ are $\mathrm{sec}^{-1}$ mole litre ${ }^{-1}$.
In the first order reaction

$$
S_{1} \xrightarrow{k_{2}} S_{2}
$$

the reaction rate is proportional to the concentration of only one reactant, namely $S_{1}$, and the rate equation is

$$
r=k_{2}\left[S_{1}\right] .
$$

The units of the reaction constant $k_{2}$ are $\sec ^{-1}$.
In the second order reaction

$$
S_{1}+S_{2} \xrightarrow{k_{3}} S_{3}
$$

the reaction rate is proportional to the product of the concentration of two reactants, namely $S_{1}$ and $S_{2}$ and the rate equation is

$$
\mathrm{r}=k_{3}\left[S_{1}\right]\left[S_{2}\right]
$$

The units of the reaction constant $k_{3}$ are $\mathrm{sec}^{-1} /\left(\right.$ mole litre $\left.^{-1}\right)$. For a reaction of order $n$, the units of reaction constant are given by $\mathrm{sec}^{-1} /\left(\text { mole litre }^{-1}\right)^{n-1}$.

### 1.3 Applications

### 1.3.1 Example - Viral Kinetics

Srivastava et al. described a simplified model of intracellular viral infection, where a single virus infects its host cell [73]. The infection process includes transcription, translation, genome replication, assembly and virus release. It is characterized by the following six reactions:

$$
\begin{align*}
& R_{1}: D N A \xrightarrow{c_{1}} m R N A  \tag{1.3a}\\
& R_{2}: m R N A \xrightarrow{c_{2}} D N A+m R N A  \tag{1.3b}\\
& R_{3}: m R N A \xrightarrow{c_{3}} \operatorname{Pr}+m R N A  \tag{1.3c}\\
& R_{4}: m R N A \xrightarrow{c_{4}} \varnothing  \tag{1.3d}\\
& R_{5}: \operatorname{Pr} \xrightarrow{c_{5}} \emptyset  \tag{1.3e}\\
& R_{6}: D N A+\operatorname{Pr} \xrightarrow{c_{6}} V r . \tag{1.3f}
\end{align*}
$$

The authors studied the viral nucleic acids, namely $D N A$ and $m R N A$ and the viral structure proteins (denoted by $\operatorname{Pr}$ in the above reactions). The $D N A$ can either transcribe into an $m R N A$ (reaction $R_{1}$ ) or it may be packaged into structural proteins to form progeny virus (denoted by $V r$ in reaction $R_{6}$ ). The $m R N A$ molecule is used as a template to duplicate the $D N A$ (reaction $R_{2}$ ) and also translate to protein (reaction $R_{3}$ ). In reactions $R_{4}$ and $R_{5}$, the $m R N A$ and the protein molecules are degraded, respectively.

### 1.3.2 Example - Gene Regulatory Network

Gene expression is a complex process consisting of transcription of a gene into an mRNA and its translation into a protein [1]. Each of the transcription and translation processes consist of a multitude of elementary reactions, and therefore cannot be modeled as instantaneous reactions. Roussel el al. and Ribeiro et al. modeled the gene expression process in a single step, using time delays [63], [65]

$$
\begin{equation*}
R N A P(t)+\operatorname{Prom}(t) \longrightarrow R N A P\left(t+\tau_{1}\right)+\operatorname{Prom}\left(t+\tau_{2}\right)+K \times \operatorname{Pr}\left(t+\tau_{3}\right) \tag{1.4}
\end{equation*}
$$

where $R N A P$ is the RNA polymerase enzyme which binds to the promotor site to initiate gene transcription, $\operatorname{Prom}$ is the promotor site, $\operatorname{Pr}$ is the resultant protein created from the translation of $m R N A, K$ is the number of copies of protein created in the process and $\tau_{1}$, $\tau_{2}$ and $\tau_{3}$ are the times required for each of the resultant products to become available. A gene regulatory network consists of transcription factors and target genes which interact with each other to determine the level of $m R N A$ and proteins created. Kierzek et al. and Ramsey et al. have performed stochastic modeling of gene regulatory networks using the SSA [44], [59].

### 1.4 Computational Models

Various modeling approaches for biochemical networks are discussed in several texts [5], [15], [78] and review papers [11], [17], [31]. The modeling approach we choose depends on the level of detail required, time scales, molecular populations, the complexity of the reaction system and so forth [5]. The deterministic Boolean models require synchronous updating and are not suitable for reactions with considerably different timescales [17]. The stochastic Petri nets are a graphical representation of the biochemical networks, and allow qualitative analysis of the reaction model [34], [61]. The most pervasive approaches are the Monte Carlo-based approaches such as the SSA and the $\tau$-leap methods to reduce the computation times of the SSA.

We will now describe the deterministic approach, the stochastic approach using the chemical master equation and the stochastic approximation such as the Fokker-Planck and the Langevin equation.

### 1.4.1 Deterministic Approach

In the deterministic approach, the state variable is the concentration of all the molecular species. These concentrations are continuous variables. From the reaction system, we obtain a system of coupled ordinary differential equations (ODEs), with one equation per molecular species. We will also refer to these as reaction rate equations (RREs).

## Example - A First Order Reaction

For the first order reaction

$$
\begin{equation*}
S_{1} \xrightarrow{k_{2}} S_{2} \tag{1.5}
\end{equation*}
$$

we have the following system of equations:

$$
\begin{gathered}
\frac{d\left[S_{1}\right]}{d t}=-k_{2}\left[S_{1}\right] \\
\frac{d\left[S_{2}\right]}{d t}=k_{2}\left[S_{1}\right] .
\end{gathered}
$$

## Example - A Second Order Reaction

For the second order reaction

$$
\begin{equation*}
S_{1}+S_{2} \xrightarrow{k_{3}} S_{3} \tag{1.6}
\end{equation*}
$$

we have the following system of equations:

$$
\begin{aligned}
\frac{d\left[S_{1}\right]}{d t} & =-k_{3}\left[S_{1}\right]\left[S_{2}\right] \\
\frac{d\left[S_{2}\right]}{d t} & =-k_{3}\left[S_{1}\right]\left[S_{2}\right] \\
\frac{d\left[S_{3}\right]}{d t} & =k_{3}\left[S_{1}\right]\left[S_{2}\right] .
\end{aligned}
$$

## Example - A Reversible First Order Reaction

For the reversible reaction

$$
S_{1} \underset{k_{2}}{\stackrel{k_{1}}{\rightleftharpoons}} S_{2}
$$

we have the following system of equations:

$$
\frac{d\left[S_{1}\right]}{d t}=-k_{1}\left[S_{1}\right]+k_{2}\left[S_{2}\right]
$$

$$
\frac{d\left[S_{2}\right]}{d t}=k_{1}\left[S_{1}\right]-k_{2}\left[S_{2}\right] .
$$

In complex systems, the above reaction rate equations, together with the initial conditions, can be solved numerically using the Euler method or the Runge-Kutta method.

In the deterministic framework we assume that the number of molecules are sufficiently high so that the discrete changes of molecular populations may be approximated by continuous changes in the concentration. However, as discussed in Section 1.1, under certain conditions, the deterministic approach will produce incorrect values for the molecular concentrations. We now give two examples comparing the deterministic method with the SSA. As mentioned earlier, SSA is a Monte Carlo-based method and it will be described in detail in Chapter 3. The stochastic parameters used in these examples are defined in Section 1.4.2. In the first example, the stochastic averages agree with the deterministic values. In the second example, the stochastic method produces a pronounced oscillatory behavior which is not predicted by the deterministic method.

## Example - Deterministic Competing Reactions

Consider the following reaction system:

$$
\begin{align*}
& R_{1}: S_{1}+S_{2} \xrightarrow{k_{1}} S_{3}  \tag{1.7a}\\
& R_{2}: S_{1}+S_{4} \xrightarrow{k_{2}} S_{5} . \tag{1.7b}
\end{align*}
$$

The RREs were solved using the Euler method. The initial concentrations were $\left[S_{1}\right]=$ 1000 molecules $\mathrm{l}^{-1},\left[S_{2}\right]=1000$ molecules $\mathrm{l}^{-1}$ and $\left[S_{4}\right]=2000$ molecules $\mathrm{l}^{-1}$. The reaction constants were $k_{1}=0.001 \mathrm{sec}^{-1} /\left(\right.$ molecules $\left.\mathrm{l}^{-1}\right), k_{2}=0.001 \mathrm{sec}^{-1} /\left(\right.$ molecules $\left.^{-1}\right)$ and plots were generated until time $t=1.0 \mathrm{sec}$. We set the volume $\mathrm{V}=1 \mathrm{l}$. The initial number of molecules were $X_{1}(t=0)=1000, X_{2}(t=0)=1000$ and $X_{4}(t=0)=2000$. The stochastic rate constants, to be defined in Section 1.4.2, were $c_{1}=0.001 \mathrm{sec}^{-1}, c_{2}=0.001 \mathrm{sec}^{-1}$. In the deterministic method, $\Delta t=0.01 \mathrm{sec}$ and in the SSA method the results were averaged over 5 realizations. Figure 1.1 shows that the simulation results using the deterministic method compare well with the averages obtained using the SSA algorithm.


Figure 1.1: In the Competing Reaction system, simulation results using the deterministic method (solid line) compare well with the averages obtained using the SSA method (dashed line).

## Example - Deterministic Lotka Reactions

In this example, we show that the results using the deterministic method do not agree with the averages obtained using the stochastic method. The Lotka reactions represent a predator-prey model proposed by Lotka in 1925 [48] and Volterra in 1930 [77] to study population dynamics

$$
\begin{align*}
& R_{1}: S_{1} \xrightarrow{k_{1}} 2 S_{1}  \tag{1.8a}\\
& R_{2}: S_{1}+S_{2} \xrightarrow{k_{2}} 2 S_{2}  \tag{1.8b}\\
& R_{3}: S_{2} \xrightarrow{k_{3}} S_{3} . \tag{1.8c}
\end{align*}
$$

In this model, the prey $S_{1}$ feeds on food with constant supply and reproduces (reaction $R_{1}$ ), and the predator $S_{2}$ feeds on the prey $S_{1}$ and reproduces (reaction $R_{2}$ ). Reaction $R_{3}$ indicates the decline of the predator due to natural causes. The reaction rate equations are

$$
\begin{align*}
& \frac{d\left[S_{1}\right]}{d t}=k_{1}\left[S_{1}\right]-k_{2}\left[S_{1}\right]\left[S_{2}\right]  \tag{1.9a}\\
& \frac{d\left[S_{2}\right]}{d t}=k_{2}\left[S_{1}\right]\left[S_{2}\right]-k_{3}\left[S_{2}\right] . \tag{1.9b}
\end{align*}
$$

The steady-state solution obtained by setting

$$
\begin{align*}
& \frac{d\left[S_{1}\right]}{d t}=0  \tag{1.10a}\\
& \frac{d\left[S_{2}\right]}{d t}=0 \tag{1.10b}
\end{align*}
$$

gives $S_{1 s s}=k_{3} / k_{2}$ and $S_{2 s s}=k_{1} / k_{2}$. That is, if we start with the initial values $S_{1}=S_{1 s s}$ and $S_{2}=S_{2 s s}$, then these values will remain constant. However, the simulations in Figure 1.2 using the SSA method show that the populations of $S_{1}$ and $S_{2}$ exhibit oscillatory behavior, with a rise in prey population followed by a rise in predator population [23]. The parameters used for these simulations were $\left[S_{1}\right]=1000$ molecules $l^{-1},\left[S_{2}\right]=1000$ molecules $\mathrm{l}^{-1}$, $k_{1}=$ $10 \mathrm{sec}^{-1} /\left(\right.$ molecules $\left.\mathrm{l}^{-1}\right), k_{2}=0.01 \mathrm{sec}^{-1} /\left(\right.$ molecules $\left.\mathrm{l}^{-1}\right)$ and $k_{3}=10 \mathrm{sec}^{-1} /\left(\right.$ molecules $\left.^{-1}\right)$.


Figure 1.2: In Lotka reactions, the SSA method produces oscillatory behavior for the predator and the prey populations, whereas the deterministic values remain constant with time.

### 1.4.2 Stochastic Approach

In the stochastic approach, we want to determine the probability that the number of molecules of the species is in a given range. We define our biochemical system as a spatially homogeneous mixture of $N$ molecular species $S_{1}, S_{2}, \ldots, S_{N}$ inside a fixed volume V. The species interact through $M$ reaction channels $R_{1}, R_{2}, \ldots R_{M}$. The state of the system is specified by the random vector $\boldsymbol{X}(t)=\left(\begin{array}{llll}X_{1}(t) & X_{2}(t) & \ldots & X_{N}(t)\end{array}\right)^{T}$, where $T$ indicates the vector transpose and $X_{i}(t), i=1,2 \ldots N$, is the number of molecules of species $S_{i}$ at time instant $t$. Let $\boldsymbol{x}(t)$ be an observation of the random vector $\boldsymbol{X}(t)$, where $\boldsymbol{x}(t)=\left(x_{1}(t) x_{2}(t)\right.$ $\left.\ldots x_{N}(t)\right)^{T}$. Define $l$ successive time indices $t_{0}<t_{1}<\ldots t_{l}$, and let $P\left(\boldsymbol{x}\left(t_{l}\right)\right)$ denote the joint probability mass function (PMF) of $\boldsymbol{X}(t)$ at any time $t_{l}$. The following argument also holds for continuous-state Markov process with $P\left(\boldsymbol{x}\left(t_{l}\right)\right)$ being the joint probability density function (PDF). Let $P_{l \mid(0,1, \ldots, l-1)}(\cdot)$ denote the conditional PMF at time $t_{l}$ given its PMF at times $t_{0}, t_{1}, \ldots, t_{l-1}$ and $P_{l \mid l-1}(\cdot)$ denote the conditional PMF at time $t_{l}$ given its PMF at time $t_{l-1}$. Then $\boldsymbol{X}(t)$ is a discrete-state continuous-time stochastic process with the Markov property

$$
\begin{align*}
& P_{l \mid(0,1, \ldots, l-1)}\left(\boldsymbol{X}\left(t_{l}\right)=\boldsymbol{x}\left(t_{l}\right) \mid \boldsymbol{X}\left(t_{0}\right)=\boldsymbol{x}\left(t_{0}\right), \boldsymbol{X}\left(t_{1}\right)=\boldsymbol{x}\left(t_{1}\right) \ldots, \boldsymbol{X}\left(t_{l-1}\right)=\boldsymbol{x}\left(t_{l-1}\right)\right) \\
& =P_{l \mid l-1}\left(\boldsymbol{X}\left(t_{l}\right)=\boldsymbol{x}\left(t_{l}\right) \mid \boldsymbol{X}\left(t_{l-1}\right)=\boldsymbol{x}\left(t_{l-1}\right)\right) \tag{1.11}
\end{align*}
$$

That is, conditional on the present state, the future state is independent of the past states. The aforementioned Markov process is completely determined by an initial state $P_{0}\left(\boldsymbol{x}\left(t_{0}\right)\right)$ and a one-step transition probability $P_{l \mid l-1}\left(\boldsymbol{X}\left(t_{l}\right)=\boldsymbol{x}\left(t_{l}\right) \mid \boldsymbol{X}\left(t_{l-1}\right)=\boldsymbol{x}\left(t_{l-1}\right)\right)$. We want to determine the time evolution of the joint probability mass function $P(\boldsymbol{x}(t))$ of $\boldsymbol{X}(t)$. (We will also denote this probability function with $P(\boldsymbol{x}, t)$ ). This is described by the master equation and its general form is given by [26], [76]:

$$
\begin{equation*}
\frac{d P(\boldsymbol{x}, t)}{d t}=\sum_{\boldsymbol{x}^{\prime}}\left[U_{\boldsymbol{x}^{\prime} \boldsymbol{x}} P\left(\boldsymbol{x}^{\prime}, t\right)-U_{\boldsymbol{x} \boldsymbol{x}^{\prime}} P(\boldsymbol{x}, t)\right], \tag{1.12}
\end{equation*}
$$

where $U_{\boldsymbol{x}^{\prime} \boldsymbol{x}}$ is the transition probability from state $\boldsymbol{x}^{\prime}$ to state $\boldsymbol{x}$ and $U_{\boldsymbol{x} \boldsymbol{x}^{\prime}}$ is the transition probability from state $\boldsymbol{x}$ to $\boldsymbol{x}^{\prime}$. It is a gain-loss equation for the probability of the state $\boldsymbol{X}=\boldsymbol{x}$.

Next, we will define reaction parameters which will give a characterization of a reaction system. Consider the following reactions:

$$
\begin{align*}
& R_{1}: \varnothing \longrightarrow S_{1}  \tag{1.13a}\\
& R_{2}: S_{1} \longrightarrow S_{2}  \tag{1.13b}\\
& R_{3}: 2 S_{1} \longrightarrow S_{2}  \tag{1.13c}\\
& R_{4}: S_{1}+S_{2} \longrightarrow S_{3} . \tag{1.13d}
\end{align*}
$$

Let $c_{m}$ be the stochastic rate constant for the reaction $R_{m}$. In the deterministic framework, the unknown quantities are the molecular concentrations and in the stochastic framework, the state variables are the number of molecules. For the four reactions in (1.13), the stochastic and the deterministic rate constants are related by:

$$
\begin{align*}
c_{1} & =k_{1} n_{A} \mathrm{~V}  \tag{1.14a}\\
c_{2} & =k_{2}  \tag{1.14b}\\
c_{3} & =k_{3} /\left(n_{A} \mathrm{~V} / 2\right)  \tag{1.14c}\\
c_{4} & =k_{4} /\left(n_{A} \mathrm{~V}\right), \tag{1.14d}
\end{align*}
$$

where $n_{A}$ is the Avogadro number. The units of $c_{m}$ are $\mathrm{sec}^{-1}$ for the four reactions. The reaction probabilities are constant for zeroth order reaction, proportional to the number of molecules for first order reactions and proportional to the product of the number of molecules for the second order reactions. Let $h_{m}(\boldsymbol{x}, t)$ be the number of distinct molecular combinations for reaction $R_{m}$ and define reaction propensity to be $a_{m}(\boldsymbol{x}, t) \equiv c_{m} h_{m}(\boldsymbol{x}, t)$. Then the probability that an $R_{m}$ reaction will take place in an infinitesimal time interval $d t$ is given by $a_{m}(\boldsymbol{x}, t) d t$, for the system in state $\boldsymbol{x}$. For the four reactions in (1.13), $h_{m}(\boldsymbol{x}, t)$ is given by:

$$
\begin{align*}
& h_{1}(\boldsymbol{x}, t)=1  \tag{1.15a}\\
& h_{2}(\boldsymbol{x}, t)=x_{1}(t)  \tag{1.15b}\\
& h_{3}(\boldsymbol{x}, t)=x_{1}(t)\left(x_{1}(t)-1\right) / 2  \tag{1.15c}\\
& h_{4}(\boldsymbol{x}, t)=x_{1}(t) x_{2}(t) . \tag{1.15d}
\end{align*}
$$

We also define $\nu_{m i}$ to be the stoichiometry of species $S_{i}$ in reaction $R_{m}$. Therefore, $\boldsymbol{\nu}_{m}=\left(\begin{array}{llll}\nu_{m 1} & \nu_{m 2} & \ldots & \nu_{m N}\end{array}\right)$ is the state-change vector for reaction $R_{m}$, and for the four reactions in (1.13), the stoichiometries are given by:

$$
\begin{align*}
& \nu_{11}=1  \tag{1.16a}\\
& \nu_{21}=-1  \tag{1.16b}\\
& \nu_{22}=1  \tag{1.16c}\\
& \nu_{31}=-2 \tag{1.16d}
\end{align*} \quad \nu_{32}=1 . \quad . \quad . \quad . \quad . \quad \nu_{41}=-1 .
$$

Using the above description, the general form of the chemical master equation (CME) is expressed as:

$$
\begin{equation*}
\frac{d P(\boldsymbol{x}, t)}{d t}=\sum_{m=1}^{M}\left[a_{m}\left(\boldsymbol{x}-\boldsymbol{\nu}_{m}, t\right) P\left(\boldsymbol{x}-\boldsymbol{\nu}_{m}, t\right)-a_{m}(\boldsymbol{x}, t) P(\boldsymbol{x}, t)\right] \tag{1.17}
\end{equation*}
$$

where $M$ is the total number of reactions in the system.

## Example - Master Equation, Competing Reactions

A derivation of the master equation is given in [25] and [76]. We will derive the master equation for the system of competing second order reactions (1.7) which we reproduce here:

$$
\begin{align*}
& R_{1}: S_{1}+S_{2} \xrightarrow{c_{1}} S_{3}  \tag{1.18a}\\
& R_{2}: S_{1}+S_{4} \xrightarrow{c_{2}} S_{5} . \tag{1.18b}
\end{align*}
$$

If we track all five species, $N=5, \boldsymbol{X}(t)=\left(X_{1}(t) X_{2}(t) X_{3}(t) X_{4}(t) X_{5}(t)\right)^{T}, P(\boldsymbol{x}, t)$ is its joint PMF, and the number of reaction channels $M=2$. Consider the time interval $(t, t+\Delta t)$ and let $o(\Delta t)$ represent terms that go to zero with $\Delta t$ faster than $\Delta t$. We use the following axioms:

- The probability of exactly one reaction $S_{1}+S_{2} \longrightarrow S_{3}$ in the time interval $(t, t+\Delta t)$ is given by $c_{1} x_{1} x_{2} \Delta t+o(\Delta t)$. The probability of exactly one reaction $S_{1}+S_{4} \longrightarrow S_{5}$ in the time interval $(t, t+\Delta t)$ is given by $c_{2} x_{1} x_{4} \Delta t+o(\Delta t)$.
- The probability that no reaction will occur in the system in the time interval $(t, t+\Delta t)$ is $\left(1-c_{1} x_{1} x_{2} \Delta t-c_{2} x_{1} x_{4} \Delta t\right)+o(\Delta t)$.
- The probability of more than one reaction in the time interval $(t, t+\Delta t)$ is $o(\Delta t)$.

Hence we obtain the following balance equation:

$$
\begin{align*}
P\left(x_{1}, x_{2}, x_{3}, x_{4}, x_{5}, t+\Delta t\right)= & \left(1-c_{1} x_{1} x_{2} \Delta t-c_{2} x_{1} x_{4} \Delta t\right) P\left(x_{1}, x_{2}, x_{3}, x_{4}, x_{5}, t\right)+ \\
& c_{1}\left(x_{1}+1\right)\left(x_{2}+1\right) \Delta t P\left(x_{1}+1, x_{2}+1, x_{3}-1, x_{4}, x_{5}, t\right)+ \\
& c_{2}\left(x_{1}+1\right)\left(x_{4}+1\right) \Delta t P\left(x_{1}+1, x_{2}, x_{3}, x_{4}+1, x_{5}-1, t\right)+ \\
& o(\Delta t) . \tag{1.19}
\end{align*}
$$

We rearrange the terms, divide by $\Delta t$, and take the limit as $\Delta t \longrightarrow 0$. All the terms of the form $o(\Delta t) / d t$ vanish and we obtain the following differential-difference equation:

$$
\begin{align*}
\frac{d P\left(x_{1}, x_{2}, x_{3}, x_{4}, x_{5}, t\right)}{d t}= & -c_{1} x_{1} x_{2} P\left(x_{1}, x_{2}, x_{3}, x_{4}, x_{5}, t\right)-c_{2} x_{1} x_{4} P\left(x_{1}, x_{2}, x_{3}, x_{4}, x_{5}, t\right)+ \\
& c_{1}\left(x_{1}+1\right)\left(x_{2}+1\right) P\left(x_{1}+1, x_{2}+1, x_{3}-1, x_{4}, x_{5}, t\right)+ \\
& c_{2}\left(x_{1}+1\right)\left(x_{4}+1\right) P\left(x_{1}+1, x_{2}, x_{3}, x_{4}+1, x_{5}-1, t\right) \tag{1.20}
\end{align*}
$$

At time $t=0$, let the number of $S_{1}, S_{2}$ and $S_{4}$ molecules be $\alpha_{1}, \alpha_{2}$ and $\alpha_{4}$, respectively. We can also describe this system with three variables, for example ( $X_{1}, X_{2}, X_{4}$ ), and rewrite the above equation

$$
\begin{align*}
\frac{d P\left(x_{1}, x_{2}, x_{4}, t\right)}{d t}= & -c_{1} x_{1} x_{2} P\left(x_{1}, x_{2}, x_{4}, t\right)-c_{2} x_{1} x_{4} P\left(x_{1}, x_{2}, x_{4}, t\right)+ \\
& c_{1}\left(x_{1}+1\right)\left(x_{2}+1\right) P\left(x_{1}+1, x_{2}+1, x_{4}, t\right)+ \\
& c_{2}\left(x_{1}+1\right)\left(x_{4}+1\right) P\left(x_{1}+1, x_{2}, x_{4}+1, t\right) \tag{1.21}
\end{align*}
$$

In Chapter 2, we will discuss various techniques for solving the master equation analytically.

### 1.4.3 Stochastic Approximation

The time evolution of a continuous stochastic process $X(t)$ is described by the integral equation [57].

$$
\begin{equation*}
X(t)=X(0)+\int_{0}^{t} b(X(s), s) d s+\int_{0}^{t} u(X(s), s) d \mathcal{W}(s) \tag{1.22}
\end{equation*}
$$

where $b(\cdot)$ is a drift coefficient, $u(\cdot)$ is a diffusion coefficient and $\mathcal{W}(\cdot)$ is a Wiener process, which is a continuous stochastic process with independent increments such that

- $E(\mathcal{W}(t)=0)$.
- For $0<t_{1}<t_{2}, \mathcal{W}\left(t_{2}\right)-\mathcal{W}\left(t_{1}\right)$ has the distribution $\mathcal{N}\left(0, t_{2}-t_{1}\right)$, where $\mathcal{N}\left(\mu, \sigma^{2}\right)$ denotes a normal random variable with mean $\mu$ and variance $\sigma^{2}$.

The $d s$-integral in (1.22) is a Reimann integral and the $d \mathcal{W}(s)$-integral is an Ito integral [45], [57]. The corresponding differential equation

$$
\begin{equation*}
d X(t)=b(\boldsymbol{X}(t), t) d t+u(\boldsymbol{X}(t), t) d \mathcal{W}(t) \tag{1.23}
\end{equation*}
$$

with the initial condition $X(t=0)=X(0)$, is known as stochastic differential equation (SDE). The Langevin equation and the Fokker-Planck equation are forms of SDE.

## Langevin Equation

The Langevin equation is a stochastic differential equation for the state of the system $\boldsymbol{X}(t)$. For the biochemical system described above, the chemical Langevin equation (CLE) is derived by Gillespie [26], [28]

$$
\begin{equation*}
d X_{i}(t)=\sum_{m=1}^{M} \nu_{m i} a_{m}(\boldsymbol{X}, t) d t+\sum_{m=1}^{M} \nu_{m i} \sqrt{a_{m}(\boldsymbol{X}, t)} d \mathcal{W}_{m i}(t) \quad i=1 \ldots N \tag{1.24}
\end{equation*}
$$

where we have set $a_{m}(\boldsymbol{x}, t) \equiv a_{m}(\boldsymbol{X}, t)$. For the above CLE to be valid, the macroscopically infinitesimal time increment $d t$ must satisfy the following two conditions:

1. $d t$ be small enough that none of the propensity functions $a_{m}(\cdot)$ changes appreciably in this time
2. $d t$ be large enough that the number of times each $R_{m}$ reaction takes place is much greater than one.

One method of solution of the Langevin equation is to use the Euler approximation and generate realizations of the stochastic process in (1.24). Several other methods of solution are discussed in Kloeden and Platen [45].

## Fokker-Planck Equation

For the Langevin equation in (1.24), there is an equivalent Fokker-Planck equation (FPE) for the PDF $P(\boldsymbol{x}, t)$ of $\boldsymbol{X}(t)[28]$

$$
\begin{align*}
\frac{\partial P(\boldsymbol{x}, t)}{\partial t}= & -\sum_{i=1}^{N} \frac{\partial}{\partial x_{i}}\left[\sum_{m=1}^{M} \nu_{m i} a_{m}(\boldsymbol{x}, t) P(\boldsymbol{x}, t)\right]+ \\
& \frac{1}{2} \sum_{i=1}^{N} \frac{\partial^{2}}{\partial x_{i}^{2}}\left[\sum_{m=1}^{M} \nu_{m i}^{2} a_{m}(\boldsymbol{x}, t) P(\boldsymbol{x}, t)\right]+ \\
& \sum_{\substack{i, j=1 \\
i<j}}^{N} \frac{\partial^{2}}{\partial x_{i} \partial x_{j}}\left[\sum_{m=1}^{M} \nu_{m i} \nu_{m j} a_{m}(\boldsymbol{x}, t) P(\boldsymbol{x}, t)\right] . \tag{1.25}
\end{align*}
$$

## Example - Fokker-Planck Equation, Competing Reactions

We track three species, $S_{1} S_{2}$ and $S_{4}, N=3$ and the number of reactions $M=2$, with reaction propensities (omitting the t's),

$$
\begin{align*}
& a_{1}(\boldsymbol{x})=c_{1} x_{1} x_{2}  \tag{1.26a}\\
& a_{2}(\boldsymbol{x})=c_{2} x_{1} x_{4} . \tag{1.26b}
\end{align*}
$$

The FPE is

$$
\begin{align*}
\frac{\partial P(\boldsymbol{x}, t)}{\partial t}= & -\frac{\partial}{\partial x_{1}}\left[(-1) c_{1} x_{1} x_{2} P(\boldsymbol{x}, t)+(-1) c_{2} x_{1} x_{4} P(\boldsymbol{x}, t)\right] \\
& -\frac{\partial}{\partial x_{2}}\left[(-1) c_{1} x_{1} x_{2} P(\boldsymbol{x}, t)\right]-\frac{\partial}{\partial x_{4}}\left[(-1) c_{2} x_{1} x_{4} P(\boldsymbol{x}, t)\right] \\
& +\frac{\partial^{2}}{\partial x_{1} \partial x_{2}}\left[c_{1} x_{1} x_{2} P(\boldsymbol{x}, t)\right]+\frac{\partial^{2}}{\partial x_{1} \partial x_{4}}\left[c_{2} x_{1} x_{4} P(\boldsymbol{x}, t)\right] \tag{1.27}
\end{align*}
$$

The Fokker-Planck equation can also be derived as an approximation of the master equation [76]. The FPE is a parabolic partial differential equation (PDE), and it is an $N$ dimensional equation for $N$ molecular species. A numerical solution of the master equation is very difficult as the state space grows exponentially with the number of variables. Elf and Sjöberg present numerical solutions of the CME and the FPE using finite difference method [14], [70]. For the master equation, the authors use an $N$-dimensional grid of size $x_{\max }$ in each dimension, with grid points at $x=0,1 \ldots x_{\max }$ and step size $\Delta x=1$. The total number of grid points is approximately $x_{\max }^{N}$. The Fokker-Planck equation is solved on a coarser grid with step size $\Delta x>1$ and the number of grid points less than $x_{\max }^{N}$. The authors compare the efficiency of the numerical solution of the Fokker-Planck equation with the SSA, and conclude that for $N>4$ the SSA would be more efficient.

## Chapter 2

## Master Equation Approach

### 2.1 Introduction

In the stochastic approach for modeling the biochemical reactions, we obtain the PMF (or PDF) of the state of the system. The problem may be addressed in several ways, including analytical solution of the master equation, Monte Carlo methods and the moment propagation methods. We give a brief review of the analytical approach.

In 1954 Rényi presented a first treatment of the stochastic analysis of the second order reaction $S_{1}+S_{2} \longrightarrow S_{3}$. He gave a recursive expression for the Laplace transform of the distribution [15], [62]. In 1958, Bartholomay derived a stochastic model of the first order reaction $S_{1} \longrightarrow S_{2}$, obtained the master equation and the distributions of the molecular populations with time, and also the limiting forms of the distributions and the standard deviation [3]. In his 1963 and 1967 papers, McQuarrie made use of the probability generating function to transform the master equation, a differential-difference equation, into a PDE, and obtained closed form expressions for a few first order reactions and the irreversible second order reaction $S_{1}+S_{2} \longrightarrow S_{3}$ [54], [55]. In 1964, Ishida studied some second order reactions and obtained expressions for the probability generating functions, which were expressed as Jacobi polynomials for the reaction $S_{1}+S_{2} \longrightarrow S_{3}$ and Legendre polynomials for the reaction $2 S_{1} \longrightarrow S_{2}$ [40]. In 2000 Laurenzi employed the Laplace transform to transform the master equation into a system of linear equations and solved the reversible second order reaction $S_{1}+S_{2} \longleftrightarrow S_{3}$ [46]. In 2005, Zhang et al. used Laplace transform to solve some first order reaction systems and obtained the molecular distributions as convolution of multinomials
[79]. In 2007, Jahnke et al. obtained solution of the CME for first order reactions. These solutions are expressed as convolution of multinomial and Poisson distributions, with timedependent parameters obtained as solutions of the reaction rate equations [41]. In this chapter, we present the results of McQuarrie and Zhang et al.

### 2.2 Method of Probability Generating Function

We review McQuarrie's results for the case of irreversible first order reactions, and give the expressions for a few other cases. We will utilize some of these results in the proposed method.

### 2.2.1 Irreversible First Order Reactions

For the first order reaction

$$
\begin{equation*}
S_{1} \xrightarrow{c_{1}} S_{2} \tag{2.1}
\end{equation*}
$$

the master equation is

$$
\begin{equation*}
\frac{d P\left(x_{1}, t\right)}{d t}=c_{1}\left(x_{1}+1\right) P\left(x_{1}+1, t\right)-c_{1} x_{1} P\left(x_{1}, t\right) \tag{2.2}
\end{equation*}
$$

where $X_{1}(t)$ is the number of molecules of $S_{1}$ at time $t$. Let $\alpha_{1}$ be the initial number of $S_{1}$ molecules. We use the probability generating function of $X_{1}(t)$

$$
\begin{equation*}
\mathcal{G}(s, t)=\sum_{x_{1}=0}^{\infty} P\left(x_{1}, t\right) s^{x_{1}}, \quad x_{1} \in \mathbb{N}_{0} \tag{2.3}
\end{equation*}
$$

where $\mathbb{N}_{0}$ is the set of natural numbers with zero. The above sum is defined for those values of $s$ for which the sum converges. Utilizing (2.3) in (2.2) yields the PDE

$$
\begin{equation*}
\frac{\partial \mathcal{G}(s, t)}{\partial t}=c_{1}(1-s) \frac{\partial \mathcal{G}(s, t)}{\partial s} \tag{2.4}
\end{equation*}
$$

with the initial condition

$$
\begin{equation*}
\mathcal{G}(s, t=0)=s^{\alpha_{1}} . \tag{2.5}
\end{equation*}
$$

We rewrite (2.4) as

$$
\begin{equation*}
\frac{\partial \mathcal{G}(s, t)}{\partial t}+c_{1}(s-1) \frac{\partial \mathcal{G}(s, t)}{\partial s}=0 \tag{2.6}
\end{equation*}
$$

This is an equation of the form

$$
\begin{equation*}
e_{1}(s, t) \frac{\partial \mathcal{G}(s, t)}{\partial t}+e_{2}(s, t) \frac{\partial \mathcal{G}(s, t)}{\partial s}=0 \tag{2.7}
\end{equation*}
$$

and its characteristic equation is [4]

$$
\begin{equation*}
\frac{d s}{d t}=\frac{e_{2}(s, t)}{e_{1}(s, t)} \tag{2.8}
\end{equation*}
$$

Its solution curves $e_{3}(s, t)=A$, where $A$ is an arbitrary constant, give the characteristic curves of the PDE. We next make a transformation from the $(s, t)$ coordinate system to the $(v, w)$ coordinate system

$$
\begin{equation*}
v=e_{3}(s, t) \quad w=s \tag{2.9}
\end{equation*}
$$

so that $w$ is constant on each curve and the PDE (2.6) becomes an ODE. The characteristic equation of (2.6) is

$$
\begin{equation*}
\frac{d s}{d t}=c_{1}(s-1) \tag{2.10}
\end{equation*}
$$

Integrating, we obtain

$$
\begin{equation*}
(s-1)=A e^{c_{1} t} \tag{2.11}
\end{equation*}
$$

The characteristic curves are given by

$$
\begin{equation*}
(s-1) e^{-c_{1} t}=A \tag{2.12}
\end{equation*}
$$

Applying the transformation of variables

$$
\begin{equation*}
v=(s-1) e^{-c_{1} t} \quad w=s \quad \mathbf{u}(v, w)=\mathcal{G}(s, t) \tag{2.13}
\end{equation*}
$$

to (2.6), we obtain

$$
\begin{equation*}
\frac{\partial \mathbf{u}}{\partial w}=0 \tag{2.14}
\end{equation*}
$$

That is, $\mathbf{u}$ is only a function of $v=(s-1) e^{-c_{1} t}$. Applying the initial condition (2.5), gives the solution

$$
\begin{equation*}
\mathcal{G}(s, t)=\left(1+(s-1) e^{-c_{1} t}\right)^{\alpha_{1}} \tag{2.15}
\end{equation*}
$$

Utilizing the expressions for the first and the second moments

$$
\begin{align*}
E\left(X_{1}(t)\right) & =\left.\frac{\partial \mathcal{G}}{\partial s}\right|_{s=1}  \tag{2.16a}\\
E\left(X_{1}^{2}(t)\right) & =\left.\frac{\partial^{2} \mathcal{G}}{\partial s^{2}}\right|_{s=1}+\left.\frac{\partial \mathcal{G}}{\partial s}\right|_{s=1} \tag{2.16b}
\end{align*}
$$

we have for the mean and the variance of $X_{1}(t)$

$$
\begin{gather*}
\mu_{1}(t) \equiv E\left(X_{1}(t)\right)=\alpha_{1} e^{-c_{1} t}  \tag{2.17}\\
\sigma_{1}^{2}(t) \equiv E\left(X_{1}^{2}(t)\right)-E^{2}\left(X_{1}(t)\right)=\alpha_{1} e^{-c_{1} t}\left(1-e^{-c_{1} t}\right) \tag{2.18}
\end{gather*}
$$

### 2.2.2 Parallel First Order Reactions

Consider the parallel first order reactions

$$
\begin{align*}
& R_{1}: S_{1} \xrightarrow{c_{1}} S_{2}  \tag{2.19a}\\
& R_{2}: S_{1} \xrightarrow{c_{2}}  \tag{2.19b}\\
& S_{3} .
\end{align*}
$$

Let $X_{1}(t)$ and $X_{2}(t)$ be the number of molecules of $S_{1}$ and $S_{2}$, respectively, and let $\alpha_{1}$ and $\alpha_{2}$ be the initial number of $S_{1}$ and $S_{2}$ molecules, respectively. For the probability generating function of $P\left(x_{1}, x_{2}, t\right)$,

$$
\begin{equation*}
\mathcal{G}\left(s_{1}, s_{2}, t\right)=\sum_{x_{1}=0}^{\infty} \sum_{x_{2}=0}^{\infty} P\left(x_{1}, x_{2}, t\right) s_{1}^{x_{1}} s_{2}^{x_{2}} \tag{2.20}
\end{equation*}
$$

we have the following expression from [55]:

$$
\begin{equation*}
\mathcal{G}\left(s_{1}, s_{2}, t\right)=\left[\frac{c_{1} s_{2}+c_{2}-\left(c_{1} s_{2}+c_{2}-C_{1} s_{1}\right) e^{-C_{1} t}}{C_{1}}\right]^{\alpha_{1}} s_{2}^{\alpha_{2}} \tag{2.21}
\end{equation*}
$$

where $C_{1}=c_{1}+c_{2}$. For the mean and the variance of $X_{1}(t)$ and $X_{2}(t)$, we have

$$
\begin{align*}
\mu_{1}(t) & =\alpha_{1} e^{-C_{1} t}  \tag{2.22}\\
\sigma_{1}^{2}(t) & =\alpha_{1} e^{-C_{1} t}\left(1-e^{-C_{1} t}\right)  \tag{2.23}\\
\mu_{2}(t) & =\alpha_{2}+\frac{c_{1} \alpha_{1}}{C_{1}}\left(1-e^{-C_{1} t}\right)  \tag{2.24}\\
\sigma_{2}^{2}(t) & =\frac{c_{1} \alpha_{1}}{C_{1}}\left(1-e^{-C_{1} t}\right)\left[1-\frac{c_{1}\left(1-e^{-C_{1} t}\right)}{C_{1}}\right] \tag{2.25}
\end{align*}
$$

### 2.2.3 Irreversible Second Order Reactions

For the second order reaction

$$
\begin{equation*}
S_{1}+S_{2} \xrightarrow{c_{1}} S_{3}, \tag{2.26}
\end{equation*}
$$

let $X_{1}(t)$ and $X_{2}(t)$ be the number of $S_{1}$ and $S_{2}$ molecules, respectively, at time $t$. Let $\alpha_{1}=X_{1}(0)$ and $\beta_{12}=X_{2}(0)-X_{1}(0)$. We track one species and the master equation is

$$
\begin{equation*}
\frac{d P\left(x_{1}, t\right)}{d t}=c_{1}\left(x_{1}+1\right)\left(\beta_{12}+x_{1}+1\right) P\left(x_{1}+1, t\right)-c_{1} x_{1}\left(\beta_{12}+x_{1}\right) P\left(x_{1}, t\right) \tag{2.27}
\end{equation*}
$$

Using the probability generating function gives the PDE

$$
\begin{equation*}
\frac{\partial \mathcal{G}(s, t)}{\partial t}=c_{1} s(1-s) \frac{\partial^{2} \mathcal{G}(s, t)}{\partial s^{2}}+c_{1}\left(\beta_{12}+1\right)(1-s) \frac{\partial \mathcal{G}(s, t)}{\partial s} \tag{2.28}
\end{equation*}
$$

This equation is solved using separation of variables and we have the following expressions for the first and second factorial moments of $X_{1}(t)$ [55]:

$$
\begin{gather*}
E\left(X_{1}(t)\right)=\sum_{n=1}^{\alpha_{1}} \frac{\left(2 n+\beta_{12}\right) \Gamma\left(\alpha_{1}+1\right) \Gamma\left(\alpha_{1}+\beta_{12}+1\right)}{\Gamma\left(\alpha_{1}-n+1\right) \Gamma\left(\alpha_{1}+\beta_{12}+n+1\right)} \Upsilon_{n}(t)  \tag{2.29}\\
E\left(X_{1}(t)\left(X_{1}(t)-1\right)\right)= \\
\sum_{n=1}^{\alpha_{1}} \frac{(n-1)\left(n+\beta_{12}+1\right)\left(2 n+\beta_{12}\right) \Gamma\left(\alpha_{1}+1\right) \Gamma\left(\alpha_{1}+\beta_{12}+1\right)}{\Gamma\left(\alpha_{1}-n+1\right) \Gamma\left(\alpha_{1}+\beta_{12}+n+1\right)} \Upsilon_{n}(t) . \tag{2.30}
\end{gather*}
$$

In the above equations, $\Gamma(\cdot)$ is the gamma function defined by

$$
\begin{equation*}
\Gamma(n)=\int_{0}^{\infty} x^{n-1} e^{-x} d x \tag{2.31}
\end{equation*}
$$

and

$$
\begin{equation*}
\Upsilon_{n}(t)=\exp \left\{-n\left(n+\beta_{12}\right) c_{1} t\right\} \tag{2.32}
\end{equation*}
$$

### 2.3 Method of Laplace Transform

In this method, Laplace transforms are employed to transform the master equation into a set of linear equations which are easier to handle.

### 2.3.1 First Order Reactions

We now present the method of Zhang et al. for computing the distributions of first order reactions [79]. This approach makes use of the fact that in a first order reaction system each molecule reacts independently of the others. Starting with one molecule from one source, the transition probabilities to other states are derived. Subsequently, the authors arrived at the distributions arising from several molecules and several sources by summing up the contribution from each. Several examples, including a chain reaction with several species, were presented.

The system has $N$ molecular species $S_{1}, S_{2}, \ldots, S_{N}, M$ reaction channels $R_{1}, R_{2}, \ldots$ $R_{M}$ and its state is specified by the random vector $\boldsymbol{X}(t)$. For the reactions

$$
\begin{equation*}
R_{m}: S_{i} \xrightarrow{c_{m}} S_{j}, \quad m=1, \ldots, M \tag{2.33}
\end{equation*}
$$

an $N \times N$ matrix D is defined as follows:

$$
\begin{align*}
& d_{i i}=-C_{i}  \tag{2.34a}\\
& d_{i j}=c_{m}, \quad i \neq j \tag{2.34b}
\end{align*}
$$

where $C_{i}$ is the sum of all the rate constants in which species $S_{i}$ is a source.

## One Source

First consider one source $S_{i}$ with one molecule. The injection of the molecule into the system is modeled by a source probability density function $f_{i}(t)$. For example, if the molecule is injected at time $t=0$, the source probability density is $f_{i}(t)=\delta(t)$, where $\delta(t)$ is the Dirac delta function. At a later time $t$, the molecule will be in any one of the states $S_{j}, j=1,2 \ldots N$, with probability $p_{i j}(t)$ satisfying the condition

$$
\begin{equation*}
\sum_{j=1}^{N} p_{i j}(t)=1 \tag{2.35}
\end{equation*}
$$

The probabilities $p_{i j}(t)$ are determined by solving the following system of first-order differential equations:

$$
\begin{equation*}
\frac{d \mathrm{p}}{d t}=\mathrm{Dp}+\mathrm{f} \tag{2.36}
\end{equation*}
$$

where

$$
\mathrm{p}=\left(\begin{array}{c}
p_{i 1}(t)  \tag{2.37}\\
p_{i 2}(t) \\
\cdot \\
\cdot \\
p_{i N}(t)
\end{array}\right), \quad \mathrm{f}=\left(\begin{array}{c}
0 \\
\cdot \\
f_{i}(t) \\
\cdot \\
0
\end{array}\right)
$$

The Laplace Transform of (2.36) gives

$$
\begin{equation*}
s \mathrm{P}=\mathrm{DP}+\mathrm{F} \tag{2.38}
\end{equation*}
$$

where $P$ and $F$ denote the Laplace transforms of $p$ and $f$, respectively. Therefore $P$ is obtained by solving

$$
\begin{equation*}
\mathrm{P}=(s I-\mathrm{D})^{-1} \mathrm{~F} . \tag{2.39}
\end{equation*}
$$

The inverse Laplace transform of P gives the required probabilities, and the joint distribution of $\left(X_{1}(t) \quad X_{2}(t) \ldots X_{N}(t)\right)$ is the categorical distribution

$$
\begin{equation*}
P\left(x_{1}, x_{2} \ldots, x_{N}, t\right)=\prod_{j=1}^{N}\left(p_{i j}(t)\right)^{x_{j}} \tag{2.40}
\end{equation*}
$$

Next, suppose that $\alpha_{i}$ molecules of $S_{i}$ are injected into the system. Since each molecule reacts independently of the others, we have $\alpha_{i}$ independent trials each with $N$ possible outcomes $X_{1}(t), X_{2}(t), \ldots$ and $X_{N}(t)$. The probabilities of these outcomes are $p_{i 1}(t), p_{i 2}(t)$, $\ldots$ and $p_{i N}(t)$, respectively, and the resulting distribution is multinomial

$$
\begin{equation*}
P\left(x_{1}, x_{2}, \ldots, x_{N}, t\right)=\frac{\alpha_{i}!}{x_{1}!x_{2}!\ldots x_{N}!}\left(p_{i 1}(t)\right)^{x_{1}}\left(p_{i 2}(t)\right)^{x_{2}} \ldots\left(p_{i N}(t)\right)^{x_{N}} \tag{2.41}
\end{equation*}
$$

where $x_{1}+x_{2}+\ldots x_{N}=\alpha_{i}$. The marginal of any one species is the binomial distribution

$$
\begin{equation*}
P\left(x_{j}, t\right)=\frac{\alpha_{i}!}{x_{j}!\left(\alpha_{i}-x_{j}\right)!}\left(p_{i j}(t)\right)^{x_{j}}\left(1-p_{i j}(t)\right)^{\alpha_{i}-x_{j}}, \quad j=1,2 \ldots N \tag{2.42}
\end{equation*}
$$

## Several Sources

If there are several independent sources, then the marginal probability density of a species is given by a convolution of binomials from each source. So for example, if we have two sources $S_{i}$ and $S_{k}$ with initial number of molecules $\alpha_{i}$ and $\alpha_{k}$, respectively, then the resulting distribution is

$$
\begin{align*}
P\left(x_{j}, t\right)= & \frac{\alpha_{i}!}{\left(x_{j}\right)!\left(\alpha_{i}-x_{j}\right)!}\left(p_{i j}(t)\right)^{x_{j}}\left(1-p_{i j}(t)\right)^{\alpha_{i}-x_{j}} \\
\otimes & \frac{\alpha_{k}!}{\left(x_{j}\right)!\left(\alpha_{k}-x_{j}\right)!}\left(p_{k j}(t)\right)^{x_{j}}\left(1-p_{k j}(t)\right)^{\alpha_{k}-x_{j}} \\
& j=1,2 \ldots N, \tag{2.43}
\end{align*}
$$

where $\sum_{j=1}^{N} p_{i j}(t)=1$ and $\sum_{j=1}^{N} p_{k j}(t)=1$. We now present some examples.

## Example - Reversible First Order Reaction

Consider the reactions

$$
\begin{align*}
& R_{1}: S_{1} \xrightarrow{c_{1}} S_{2}  \tag{2.44a}\\
& R_{2}: S_{2} \xrightarrow{c_{2}} \quad S_{1} \tag{2.44b}
\end{align*}
$$

with one source $S_{1}$ consisting of $\alpha_{1}$ molecules. For each $S_{1}$ molecule, we have the following system of equations:

$$
\binom{\frac{d p_{11}(t)}{d t}}{\frac{d p_{12}(t)}{d t}}=\left(\begin{array}{cc}
-c_{1} & c_{2}  \tag{2.45}\\
c_{1} & -c_{2}
\end{array}\right)\binom{p_{11}(t)}{p_{12}(t)}+\binom{\delta(t)}{0} .
$$

Taking the Laplace transform of the above equation, and solving for $P$ gives

$$
\begin{equation*}
\mathrm{P}=\binom{\frac{s+c_{2}}{s\left(s+c_{1}+c_{2}\right)}}{\frac{c_{1}}{s\left(s+c_{1}+c_{2}\right)}} \tag{2.46}
\end{equation*}
$$

The inverse Laplace transform of the above expression yields the following one molecule probabilities:

$$
\begin{align*}
& p_{11}(t)=\frac{1}{c_{1}+c_{2}}\left(c_{2}+c_{1} e^{-\left(c_{1}+c_{2}\right) t}\right)  \tag{2.47a}\\
& p_{12}(t)=\frac{c_{1}}{c_{1}+c_{2}}\left(1-e^{-\left(c_{1}+c_{2}\right) t}\right) . \tag{2.47b}
\end{align*}
$$

With $\alpha_{1}$ initial molecules of $S_{1}$, the distributions of $X_{1}(t)$ and $X_{2}(t)$ are given by

$$
\begin{equation*}
P\left(x_{j}, t\right)=\frac{\alpha_{1}!}{x_{j}!\left(\alpha_{1}-x_{j}\right)!}\left(p_{1 j}(t)\right)^{x_{j}}\left(1-p_{1 j}(t)\right)^{\alpha_{1}-x_{j}}, \quad j=1,2 . \tag{2.48}
\end{equation*}
$$

## Example - Three Species

Consider the reactions

$$
\begin{align*}
& R_{1}: S_{1} \xrightarrow{c_{1}} S_{2}  \tag{2.49a}\\
& R_{2}: S_{2} \xrightarrow{c_{2}} S_{3}, \tag{2.49b}
\end{align*}
$$

with one source $S_{1}$ consisting of $\alpha_{1}$ molecules. Once again, for each $S_{1}$ molecule, we have the following system of equations:

$$
\left(\begin{array}{c}
\frac{d p_{11}(t)}{d t}  \tag{2.50}\\
\frac{d p_{12}(t)}{d t} \\
\frac{d p_{13}(t)}{d t}
\end{array}\right)=\left(\begin{array}{ccc}
-c_{1} & 0 & 0 \\
c_{1} & -c_{2} & 0 \\
0 & c_{2} & 0
\end{array}\right)\left(\begin{array}{c}
p_{11}(t) \\
p_{12}(t) \\
p_{13}(t)
\end{array}\right)+\left(\begin{array}{c}
\delta(t) \\
0 \\
0
\end{array}\right)
$$

Using the above method yields the following one molecule probabilities:
Case $i \quad c_{1} \neq c_{2}$

$$
\begin{align*}
& p_{11}(t)=e^{-c_{1} t}  \tag{2.51a}\\
& p_{12}(t)=\frac{c_{1}}{c_{2}-c_{1}}\left(e^{-c_{1} t}-e^{-c_{2} t}\right)  \tag{2.51~b}\\
& p_{13}(t)=1-\frac{c_{2}}{c_{2}-c_{1}} e^{-c_{1} t}+\frac{c_{1}}{c_{2}-c_{1}} e^{-c_{2} t} . \tag{2.51c}
\end{align*}
$$

Case ii $c_{1}=c_{2}$

$$
\begin{align*}
& p_{11}(t)=e^{-c_{1} t}  \tag{2.52a}\\
& p_{12}(t)=c_{1} t e^{-c_{1} t}  \tag{2.52b}\\
& p_{13}(t)=1-e^{-c_{1} t}-c_{1} t e^{-c_{1} t} \tag{2.52c}
\end{align*}
$$

With $\alpha_{1}$ initial molecules of $S_{1}$, the distributions of $X_{1}(t), X_{2}(t)$ and $X_{3}(t)$ are given by (2.42), with $i=1$ and $N=3$.

## Chapter 3

## Monte Carlo Methods

### 3.1 Introduction

The SSA method generates complete realizations of the Markov process $\boldsymbol{X}(t)$, and it was formulated for biochemical reaction systems by Gillespie in 1976 [22], [23]. A form of this method was introduced by Doob in 1945 [12]. In this method, the system evolves one reaction at a time, and uses two samples, one each, to determine the time to next reaction and the index of next reaction, respectively. The marginal densities and other statistics are inferred from the generated realizations. However, simulating only one reaction per time step is a major drawback of the SSA algorithm. If some molecular population is very large, the time to next reaction is very small and this renders the algorithm too slow to be useful. There are two different formulations of the SSA, namely, the direct method and the first reaction method. Gibson et al. modified the first reaction method to produce the next reaction method. In this method, the time to next reaction is sampled separately for each reaction, and the next reaction to be executed is the one with the smallest such time [20]. All of the samples are utilized, and this method scales better than SSA with the number of reactions. Cao et al. proposed an optimization of the direct method and compare it with the next reaction method [7]. Both the SSA and the next reaction methods are exact, as the trajectories generated by these methods would be the same as those generated from the solution of the master equation.

Numerous approximate methods have been established to reduce the computation time. In the $\tau$-leap algorithm introduced by Gillespie, one steps through time in intervals large
enough to contain several firings of each reaction channel [27]. Each reaction channel fires independently of the others, with the number of such firings exhibiting Poisson distribution. However, the range of Poisson random variable varies from zero to infinity, and can cause negative molecular populations in the $\tau$-leap method. Tian et al. [75] and Chatterjee et al. [9] proposed $\tau$-leap methods in which the reaction channels fire from a binomial distribution. Pettigrew et al. developed a multinomial $\tau$-leap method, which is an extension of the binomial $\tau$-leap method, and increases the speed by partitioning a large reaction network into smaller groups [58]. Slepoy et al. proposed a method to reduce the computation time for very large reaction networks [72]. They use a method called composition and rejection, for faster computation of the index to next reaction. Their algorithm assumes that the average number of coupling between reactions does not grow as the size of the reaction network grows. In this chapter, we will describe the SSA and the $\tau$-leap algorithm proposed by Tian et al.

### 3.2 Stochastic Simulation Algorithm

In this section, we recount the stochastic simulation algorithm [22], [23]. We have a system of $N$ molecular species $S_{1}, S_{2}, \ldots, S_{N}$ distributed uniformly inside a fixed volume V. The number of non-reactive molecular collisions occur much more frequently than reactive molecular collisions. The mixture is in thermal equilibrium, so that the collisions occur randomly. The species interact through $M$ reaction channels $R_{1}, R_{2}, \ldots R_{M}$. The state of the system is specified by the random vector $\boldsymbol{X}(t)$. As described in Section 1.4.2, $h_{m}(\boldsymbol{x}, t)$ is the number of distinct molecular combinations for reaction $R_{m}$ and the reaction propensity is defined by $a_{m}(\boldsymbol{x}, t) \equiv c_{m} h_{m}(\boldsymbol{x}, t)$. Then the probability that an $R_{m}$ reaction will take place in an infinitesimal time interval $d t$ is given by $a_{m}(\boldsymbol{x}, t) d t$, for the system in state $\boldsymbol{x}$. The algorithm then generates samples of the following events:

- The time to next reaction.
- The index of next reaction.

Let $P(\tau, k)$ denote the next reaction probability density, where $\tau$ is the time to next reaction and $k$ is the index of next reaction. Then $P(\tau, k) d \tau$ gives the probability that the
next reaction will occur in time $d \tau$, and it will be the reaction $R_{k} . P(\tau, k)$ takes the form

$$
\begin{equation*}
P(\tau, k)=a_{k}(\boldsymbol{x}, t) e^{-a_{0}(\boldsymbol{x}, t) \tau}, \quad \tau>0, \quad k=1,2, \ldots M \tag{3.1}
\end{equation*}
$$

where

$$
\begin{equation*}
a_{0}(\boldsymbol{x}, t)=\sum_{m=1}^{M} a_{m}(\boldsymbol{x}, t) \tag{3.2}
\end{equation*}
$$

Marginalizing $P(\tau, k)$, with respect to $k$, gives the PDF of $\tau$

$$
\begin{equation*}
P_{1}(\tau)=a_{0}(\boldsymbol{x}, t) e^{-a_{0}(\boldsymbol{x}, t) \tau}, \quad \tau>0 \tag{3.3}
\end{equation*}
$$

an exponential distribution with parameter $a_{0}(\boldsymbol{x}, t)$. Again, marginalizing $P(\tau, k)$, with respect to $\tau$, gives the index of next reaction with probability

$$
\begin{equation*}
\operatorname{Prob}(k=m)=\frac{a_{m}(\boldsymbol{x}, t)}{a_{0}(\boldsymbol{x}, t)} . \tag{3.4}
\end{equation*}
$$

Therefore $\tau$ is sampled from the exponential distribution (3.3) and the index $k$ is sampled from the categorical distribution, with $M$ possible outcomes, and probabilities $a_{1} / a_{0}, a_{2} / a_{0}$ $\ldots, a_{M} / a_{0}$.

### 3.2.1 SSA Implementation

Given the reaction stoichiometries, reaction constants, the initial number of molecules $\boldsymbol{X}(t=0)$, the stopping time $T_{s}$ and the number of simulations, the algorithm is implemented as follows:
For each simulation, perform the following initialization and recursion:

## Initialization

- Set the initial time $t=0$ and the molecular populations to the given initial values.
- Compute the values of $h_{m}(\boldsymbol{x}, t)$ and $a_{m}(\boldsymbol{x}, t), m=1 \ldots M$.
- Compute the value of $a_{0}(\boldsymbol{x}, t)$.


## Recursion

Repeat the following steps until stop time $t=T_{s}$ is reached:

- Generate a random number $\tau$ from distribution $P_{1}(\tau)$.
- Generate a random number $m$ from the multinomial distribution with probabilities $a_{1} / a_{0}, a_{2} / a_{0} \ldots, a_{M} / a_{0}$. This gives the reaction $R_{m}$ to be executed.
- Update the molecular populations of all the species affected by the reaction $R_{m}$.
- Update $a_{m}(\boldsymbol{x}, t)$ and $a_{0}(\boldsymbol{x}, t)$.
- Advance the time to $t+\tau$.

Each simulation of the above algorithm generates a realization of the biochemical process from time $t=0$ to time $t=T_{s}$. The marginal densities $P_{1}\left(x_{1}, t\right), P_{2}\left(x_{2}, t\right), \ldots, P_{N}\left(x_{N}, t\right)$, moments of $\boldsymbol{X}(t)$ and other statistics can be obtained from these realizations.

### 3.3 Binomial $\tau$-leap Method

In the $\tau$-leap algorithms, the time is divided into contiguous time intervals, and the system leaps along the time axis, to allow for many reaction events to take place in these time intervals [27]. We will now recite the $\tau$-leap algorithm proposed by Tian et al. [75]. Once again, we have a uniform mixture of $N$ molecular species $S_{1}, S_{2}, \ldots, S_{N}$, which interact through $M$ reaction channels $R_{1}, R_{2}, \ldots R_{M}$. Define a parameter $K_{m}$ for the first and second order reactions as follows:

$$
\begin{align*}
& R_{m}: S_{1} \longrightarrow S_{2}, \quad K_{m}=x_{1}(t)  \tag{3.5a}\\
& R_{m}: 2 S_{1} \longrightarrow S_{2}, \quad K_{m}=\left\lfloor\frac{x_{1}(t)}{2}\right\rfloor  \tag{3.5b}\\
& R_{m}: S_{1}+S_{2} \longrightarrow S_{3}, \quad K_{m}=\min \left(x_{1}(t), x_{2}(t)\right), \tag{3.5c}
\end{align*}
$$

where $\lfloor x\rfloor$ denotes largest integer less than or equal to $x$.
A time step $\tau$ is selected, and the criteria for doing so will be described shortly. At each time step, and for each reaction channel $R_{m}, m=1 \ldots M$, we draw a sample $L_{m}$ from the
distribution

$$
\begin{equation*}
\operatorname{bin}\left(K_{m}, \frac{a_{m}(\boldsymbol{x}, t) \tau}{K_{m}}\right) \tag{3.6}
\end{equation*}
$$

where $\operatorname{bin}(K, p)$ denotes binomial distribution with $K$ independent trials each of which is successful with probability $p$. After sampling from all the channels, the molecular populations are updated as follows:

$$
\begin{equation*}
\boldsymbol{x}(t+\tau)=\boldsymbol{x}(t)+\sum_{m=1}^{M} \boldsymbol{\nu}_{m} L_{m} . \tag{3.7}
\end{equation*}
$$

The time step $\tau$ must satisfy the condition,

$$
\begin{equation*}
0 \leq \frac{a_{m}(\boldsymbol{x}, t) \tau}{K_{m}} \leq 1 \tag{3.8}
\end{equation*}
$$

In addition, $\tau$ should be consistent with the leap condition which states that no propensity function must change appreciably in this time [29]. The leap condition is mathematically formulated as

$$
\begin{equation*}
\left|a_{m}(\boldsymbol{x}, t+\tau)-a_{m}(\boldsymbol{x}, t)\right| \leq \epsilon a_{0}(\boldsymbol{x}, t), \quad m=1, \ldots M \tag{3.9}
\end{equation*}
$$

where $a_{0}(\boldsymbol{x}, t)$ is defined in (3.2) and $\epsilon$ is some specified error control parameter. Starting from the above leap condition, $\tau$ is found to be

$$
\begin{equation*}
\tau=\min _{m \in[1, M]}\left(\frac{\epsilon a_{0}(\boldsymbol{x}, t)}{\left|\mu_{\tau_{m}}(\boldsymbol{x}, t)\right|}, \frac{\epsilon^{2} a_{0}^{2}(\boldsymbol{x}, t)}{\sigma_{\tau_{m}}^{2}(\boldsymbol{x}, t)}\right), \tag{3.10}
\end{equation*}
$$

where

$$
\begin{array}{ll}
\mu_{\tau_{m}}(\boldsymbol{x}, t)=\sum_{m^{\prime}=1}^{M} \eta_{m m^{\prime}}(\boldsymbol{x}, t) a_{m^{\prime}}(\boldsymbol{x}, t) & m=1, \ldots, M \\
\sigma_{\tau_{m}}^{2}(\boldsymbol{x}, t)=\sum_{m^{\prime}=1}^{M} \eta_{m m^{\prime}}^{2}(\boldsymbol{x}, t) a_{m^{\prime}}(\boldsymbol{x}, t) & m=1, \ldots, M \\
\eta_{m m^{\prime}}(\boldsymbol{x}, t)=\sum_{i=1}^{N} \frac{\partial a_{m}(\boldsymbol{x}, t)}{\partial x_{i}} v_{m^{\prime} i} & m, m^{\prime}=1, \ldots, M \tag{3.11c}
\end{array}
$$

### 3.3.1 Simultaneous Sampling

In the above method, if the same species $S_{i}$ undergoes reactions in two reaction channels $R_{m}$ and $R_{m^{\prime}}$, then sample as follows:

- Determine $K_{m}$ and $K_{m^{\prime}}$ for the reactions $R_{m}$ and $R_{m^{\prime}}$, respectively, from equations (3.5).
- Generate a sample $L_{m m^{\prime}}$ for the total number of reactions that $S_{i}$ undergoes in channels $R_{m}$ and $R_{m^{\prime}}$ That is sample from the distribution

$$
\begin{equation*}
\operatorname{bin}\left(K_{i}, \frac{\left(a_{m}(\boldsymbol{x}, t)+a_{m^{\prime}}(\boldsymbol{x}, t)\right) \tau}{K_{i}}\right) \tag{3.12}
\end{equation*}
$$

under the conditions

$$
\begin{equation*}
K_{i}=\min \left(K_{m}, K_{m^{\prime}}\right) \neq 0 \tag{3.13}
\end{equation*}
$$

and

$$
\begin{equation*}
0 \leq \frac{\left(a_{m}(\boldsymbol{x}, t)+a_{m^{\prime}}(\boldsymbol{x}, t)\right) \tau}{K_{i}} \leq 1 \tag{3.14}
\end{equation*}
$$

- Next generate a sample value $L_{m}$ for the total number of reactions that $S_{i}$ undergoes in channel $R_{m}$. That is sample from

$$
\begin{equation*}
\operatorname{bin}\left(L_{m m^{\prime}}, \frac{a_{m}(\boldsymbol{x}, t)}{a_{m}(\boldsymbol{x}, t)+a_{m^{\prime}}(\boldsymbol{x}, t)}\right) . \tag{3.15}
\end{equation*}
$$

- The number of reactions in channel $R_{m^{\prime}}$ is $L_{m^{\prime}}=L_{m m^{\prime}}-L_{m}$.

The authors give an example with twenty two reactions, where the $\tau$-leap method shows a gain in speed by a factor of fifteen, compared with the SSA for $\epsilon=0.01$. Since the SSA is used as a benchmark by most authors in the accelerated Monte Carlo-based and other methods, we will compare our proposed method with the SSA for both accuracy and speed.

## Chapter 4

## Proposed Method - First Order Reactions

### 4.1 Introduction

We propose a method for computing recursively, with time, the first two moments of the state of the system, and we call it the recursive moment (RM) method [50], [51]. The derivation of our method is based on conditional expectation relations of the random variables, Markovian properties of the biochemical system and approximations for small time steps. Our goal is to derive general expressions for very complex systems consisting of a large number of species and reactions. A second goal is that the expressions be straightforward to implement.

General expressions for the first order reactions are presented in this chapter and for the second order reactions are presented in the next chapter. The recipe for the first order reactions is simpler and can be used for a system consisting of first order reactions and those second order reactions which may be approximated as first order. The recipe for second order reactions can be used for systems consisting of both the first and the second order reactions. We compare the RM method with the SSA for accuracy, and further comparisons are characterized in Chapter 6.

### 4.2 Recursive Moment Method

Recall that we have a spatially homogeneous mixture of $N$ molecular species $S_{1}, S_{2}, \ldots$, $S_{N}$ inside a fixed volume V . The species interact through $M$ reaction channels $R_{1}, R_{2}, \ldots$ $R_{M}$. The state of the system is specified by the random vector $\boldsymbol{X}(t)=\left(X_{1}(t) \quad X_{2}(t) \ldots\right.$ $\left.X_{N}(t)\right)^{T}$, where $X_{i}(t), i=1,2 \ldots N$, is the number of molecules of species $S_{i}$ at time instant $t$.

The joint non-central moments of $\boldsymbol{X}(t)$ are given by [8], [74]

$$
\begin{equation*}
\gamma^{\boldsymbol{q}}(t) \equiv \gamma^{\left(q_{1} q_{2} \ldots q_{N}\right)}(t)=E\left(X_{1}^{q_{1}}(t) X_{2}^{q_{2}}(t) \ldots X_{N}^{q_{N}}(t)\right), \tag{4.1}
\end{equation*}
$$

where $\boldsymbol{q}=\left(q_{1} q_{2} \ldots q_{N}\right) \in \mathbb{N}^{N}$ and $\mathbb{N}$ is the set of positive integers. We will also denote $\left(X_{1}^{q_{1}}(t) X_{2}^{q_{2}}(t) \ldots X_{N}^{q_{N}}(t)\right)$ by $\boldsymbol{X}^{\boldsymbol{q}}(t)$. The joint central moments of $\boldsymbol{X}(t)$ are given by

$$
\begin{align*}
& \gamma^{\prime \boldsymbol{q}}(t) \equiv \gamma^{\prime\left(q_{1} q_{2} \ldots q_{N}\right)}(t) \\
& =E\left(\left(X_{1}(t)-\mu_{1}(t)\right)^{q_{1}}\left(X_{2}(t)-\mu_{2}(t)\right)^{q_{2}} \ldots\left(X_{N}(t)-\mu_{N}(t)\right)^{q_{N}}(t)\right) \tag{4.2}
\end{align*}
$$

where

$$
\begin{equation*}
\mu_{i}(t)=E\left(X_{i}(t)\right), \quad i=1, \ldots N . \tag{4.3}
\end{equation*}
$$

The order of the moments $Q$ is defined by $Q=\sum_{n=1}^{N} q_{n}$.
Further, let $\boldsymbol{\mu}(t)$ denote the mean of $\boldsymbol{X}(t)$ with elements $\mu_{i}(t), i=1,2 \ldots N$, and $\boldsymbol{C}(t)$ denote the covariance matrix of $\boldsymbol{X}(t)$ with its elements given by

$$
\begin{equation*}
\sigma_{i j}(t)=E\left(\left(X_{i}(t)-\mu_{i}(t)\right)\left(X_{j}(t)-\mu_{j}(t)\right)\right), \quad i, j=1, \ldots N . \tag{4.4}
\end{equation*}
$$

The diagonal elements of $\boldsymbol{C}(t)$ will be denoted by $\sigma_{i}^{2}(t)$. In the proposed method, we will compute the mean and the covariance matrix of $\boldsymbol{X}(t)$.

In this chapter, we will discuss reactions of the type

$$
\begin{equation*}
R_{m}: S_{i} \xrightarrow{c_{m}} S_{j}, \quad m=1, \ldots, M \tag{4.5}
\end{equation*}
$$

consisting of one reactant and one product. For a system consisting of one reaction as in (4.5), (2.15) shows that the distribution of $X_{i}$ is binomial, where the probability that a
molecule of species $S_{i}$ converts to a molecule of species $S_{j}$ in time $\Delta t$ is $1-e^{-c_{m} \Delta t}$, and the probability that it stays as a molecule of species $S_{i}$ is $e^{-c_{m} \Delta t}$.

We choose a time step $\Delta t$ between two recursions and the criterion for doing so is discussed in Section 4.3. An $N \times N$ matrix of transition probabilities $\boldsymbol{P}$ is defined such that its $i j$-th element, $p_{i j}$, represents the probability that a molecule of species $S_{i}$ converts to a molecule of species $S_{j}$ during $\Delta t$. These elements are given by

$$
\begin{align*}
& p_{i i}=e^{-C_{i} \Delta t}  \tag{4.6a}\\
& p_{i j}=\frac{c_{m}}{C_{i}}\left(1-e^{-C_{i} \Delta t}\right), \quad i \neq j, \tag{4.6b}
\end{align*}
$$

where $C_{i}$ is the sum of all the rate constants in which species $S_{i}$ is a source. These probabilities are assumed constant during $\Delta t$.

For each molecular species in the system, and at each time step, we obtain expressions for the first two moments conditioned on the previous time step. We employ the following relations for conditional expectations [8], [16]:

$$
\begin{align*}
E_{U}(U) & =E_{Z}\left(E_{U \mid Z}(U \mid Z)\right)  \tag{4.7}\\
E_{U \mid Z}\left(U^{2} \mid Z\right) & =\operatorname{var}(U \mid Z)+E_{U \mid Z}^{2}(U \mid Z)  \tag{4.8}\\
E_{U, V \mid Z}(U V \mid Z) & =\operatorname{cov}(U, V \mid Z)+E_{U \mid Z}(U \mid Z) E_{V \mid Z}(V \mid Z), \tag{4.9}
\end{align*}
$$

where $U, V$ and $Z$ are random variables, $E(\cdot)$ is the expectation operator, $\operatorname{var}(\cdot)$ is the variance of a random variable and $\operatorname{cov}(\cdot)$ indicates the covariance between two random variables. The subscript of $E$ indicates that the sums (or integrals) are evaluated with respect to the probability mass function (or probability density function) of that random variable. These subscripts will be omitted whenever it is clear from the context how the integrals are evaluated. For example, $E_{U \mid Z}(U \mid Z)$ in the above expression is the conditional expectation of $U$ given $Z$. Marginalizing (4.8) and (4.9) gives the required second moments

$$
\begin{align*}
E_{U}\left(U^{2}\right) & =E_{Z}\left(E_{U \mid Z}\left(U^{2} \mid Z\right)\right)  \tag{4.10}\\
E_{U, V}(U V) & =E_{Z}\left(E_{U, V \mid Z}(U V \mid Z)\right) \tag{4.11}
\end{align*}
$$

We also define covariance and conditional covariance of the random vectors $\boldsymbol{U}, \boldsymbol{V}$ and $\boldsymbol{Z}$ as follows:

$$
\begin{align*}
\operatorname{cov}(\boldsymbol{U}, \boldsymbol{V}) & =E(\boldsymbol{U} \boldsymbol{V})-E(\boldsymbol{U}) E(\boldsymbol{V})  \tag{4.12}\\
\operatorname{cov}(\boldsymbol{U}, \boldsymbol{V} \mid \boldsymbol{Z}) & =E(\boldsymbol{U} \boldsymbol{V} \mid \boldsymbol{Z})-E(\boldsymbol{U} \mid \boldsymbol{Z}) E(\boldsymbol{V} \mid \boldsymbol{Z}) \tag{4.13}
\end{align*}
$$

For the random vector $\boldsymbol{X}(t+\Delta t)$, we have from relations (4.7) - (4.9)

$$
\begin{align*}
& \boldsymbol{\mu}(t+\Delta t) \equiv E_{\boldsymbol{X}(t+\Delta t)}(\boldsymbol{X}(t+\Delta t)) \\
& =E_{\boldsymbol{X}(t)}\left(E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t))\right)  \tag{4.14}\\
& E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}\left(\boldsymbol{X}(t+\Delta t) \boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \\
& =\operatorname{cov}\left(\boldsymbol{X}(t+\Delta t), \boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \\
& \quad+E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)) E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}\left(\boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \tag{4.15}
\end{align*}
$$

To evaluate (4.14), we obtain the expression for $E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t))$ by summing up the contributions from all the species. This yields

$$
\begin{equation*}
E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t))=\boldsymbol{P}^{T} \boldsymbol{X}(t) \tag{4.16}
\end{equation*}
$$

and after taking the expectation $E_{\boldsymbol{X}(t)}(\cdot)$ of the above expression, we obtain

$$
\begin{equation*}
\boldsymbol{\mu}(t+\Delta t)=\boldsymbol{P}^{T} \boldsymbol{\mu}(t) \tag{4.17}
\end{equation*}
$$

We next evaluate (4.15). The second term is

$$
\begin{align*}
& E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)) E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}\left(\boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \\
& =\boldsymbol{P}^{T} \boldsymbol{X}(t) \boldsymbol{X}^{T}(t) \boldsymbol{P} \tag{4.18}
\end{align*}
$$

where we have utilized (4.16). To evaluate the first term in (4.15), we set

$$
\begin{equation*}
\boldsymbol{R}(t) \equiv \operatorname{cov}\left(\boldsymbol{X}(t+\Delta t), \boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) . \tag{4.19}
\end{equation*}
$$

The elements of $\boldsymbol{R}(t)$ are obtained by utilizing the multinomial distribution approximation as follows. Given $K_{m}$ molecules of species $S_{m}$ at time $t$, let $p_{m 1}, p_{m 2}, \ldots, p_{m N}$ be the probabilities that species $S_{m}$ will become species $S_{1}, S_{2} \ldots, S_{N}$, respectively. Then the joint probability of the number of molecules $\left(k_{m 1}, k_{m 2}, \ldots, k_{m N}\right)$ of species $S_{1}, S_{2}, \ldots S_{N}$, respectively, created by the $K_{m}$ molecules of $S_{m}$ during the interval $(t, t+\Delta t)$ is given by the multinomial distribution

$$
\begin{equation*}
f\left(k_{m 1}, k_{m 2}, \ldots, k_{m N}\right)=\frac{K_{m}!}{k_{m 1}!k_{m 2}!\ldots k_{m N}!} p_{m 1}^{k_{m 1}} p_{m 2}^{k_{m 2}} \ldots p_{m N}^{k_{m N}} \tag{4.20}
\end{equation*}
$$

where $\sum_{i=1}^{N} k_{m i}=K_{m}$ and $\sum_{i=1}^{N} p_{m i}=1$. From the moment generating function of the multinomial distribution

$$
\begin{equation*}
\mathcal{M}_{\boldsymbol{X}(t)}\left(\omega_{1}, \omega_{2}, \ldots, \omega_{N-1}\right)=\left(\sum_{i=1}^{N-1} p_{m i} e^{\omega_{i}}+p_{m N}\right)^{K_{m}} \tag{4.21}
\end{equation*}
$$

we compute the expectations $E\left(k_{m i}^{2}\right)$ and $E\left(k_{m i} k_{m j}\right)$ and readily obtain

$$
\begin{gather*}
\operatorname{var}\left(k_{m i}\right)=K_{m} p_{m i}\left(1-p_{m i}\right)  \tag{4.22}\\
\operatorname{cov}\left(k_{m i}, k_{m j}\right)=-K_{m} p_{m i} p_{m j} \tag{4.23}
\end{gather*}
$$

We can also arrive at the above result by noting that the marginal of any random variable in (4.20) is binomial and the marginal of any two random variables in (4.20) is trinomial.

We assume that in the time interval $(t, t+\Delta t)$, the distribution of the population of each reactant species evolves as multinomial, and contributes to all of its product species. Using (4.22) and (4.23) and summing up the contributions from all the species yields the elements of $\boldsymbol{R}(t)$

$$
\begin{align*}
r_{i i}(t) & =\sum_{k=1}^{N} X_{k}(t) p_{k i}\left(1-p_{k i}\right)  \tag{4.24a}\\
r_{i j}(t) & =-\sum_{k=1}^{N} X_{k}(t) p_{k i} p_{k j} \tag{4.24b}
\end{align*} \quad i \neq j .
$$

We will also need the expectation $\boldsymbol{B}(t) \equiv E_{\boldsymbol{X}(t)}(\boldsymbol{R}(t))$. Its elements are given by

$$
\begin{align*}
& b_{i i}(t)=\sum_{k=1}^{N} \mu_{k}(t) p_{k i}\left(1-p_{k i}\right)  \tag{4.25a}\\
& b_{i j}(t)=-\sum_{k=1}^{N} \mu_{k}(t) p_{k i} p_{k j} \quad i \neq j, \tag{4.25b}
\end{align*}
$$

and $\mu_{j}(t)=E\left(X_{j}(t)\right)$. Substituting (4.18) and (4.19) in (4.15) yields

$$
\begin{equation*}
E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}\left(\boldsymbol{X}(t+\Delta t) \boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right)=\boldsymbol{R}(t)+\boldsymbol{P}^{T} \boldsymbol{X}(t) \boldsymbol{X}^{T}(t) \boldsymbol{P} \tag{4.26}
\end{equation*}
$$

Marginalizing the above expression gives

$$
\begin{equation*}
E\left(\boldsymbol{X}(t+\Delta t) \boldsymbol{X}^{T}(t+\Delta t)\right)=\boldsymbol{B}(t)+\boldsymbol{P}^{T} E_{\boldsymbol{X}(t)}\left(\boldsymbol{X}(t) \boldsymbol{X}^{T}(t)\right) \boldsymbol{P} \tag{4.27}
\end{equation*}
$$

The covariance matrix is given by

$$
\begin{align*}
\boldsymbol{C}(t+\Delta t) & =E\left(\boldsymbol{X}(t+\Delta t) \boldsymbol{X}^{T}(t+\Delta t)\right)-\boldsymbol{\mu}(t+\Delta t) \boldsymbol{\mu}^{T}(t+\Delta t) \\
& =\boldsymbol{B}(t)+\boldsymbol{P}^{T} E_{\boldsymbol{X}(t)}\left(\boldsymbol{X}(t) \boldsymbol{X}^{T}(t)\right) \boldsymbol{P}-\boldsymbol{P}^{T} \boldsymbol{\mu}(t) \boldsymbol{\mu}^{T}(t) \boldsymbol{P} \tag{4.28}
\end{align*}
$$

where we have utilized (4.17) and (4.27) in the last equality. Therefore, the recursive expression for the covariance matrix is

$$
\begin{equation*}
\boldsymbol{C}(t+\Delta t)=\boldsymbol{B}(t)+\boldsymbol{P}^{T} \boldsymbol{C}(t) \boldsymbol{P} . \tag{4.29}
\end{equation*}
$$

### 4.2.1 Example - A Simple Reversible Reaction

Consider the following system of reactions:

$$
\begin{align*}
& R_{1}: S_{1} \xrightarrow{c_{1}} S_{2}  \tag{4.30a}\\
& R_{2}: S_{2} \xrightarrow{c_{2}} S_{1} . \tag{4.30b}
\end{align*}
$$

For reaction $R_{1}$, the expressions for the mean and the variance of $X_{1}(t)$ due to McQuarrie are [55]

$$
\begin{gather*}
E\left(X_{1}(t)\right)=X_{1}(t=0) e^{-c_{1} t}  \tag{4.31}\\
\operatorname{var}\left(X_{1}(t)\right)=X_{1}(t=0) e^{-c_{1} t}\left(1-e^{-c_{1} t}\right) . \tag{4.32}
\end{gather*}
$$

We have analogous expressions for the mean and the variance of $X_{2}(t)$. By summing up the contributions from reactions $R_{1}$ and $R_{2}$, in time interval $\Delta t$, we have the following expressions for the conditional expectations of $X_{1}(t+\Delta t)$ and $X_{1}^{2}(t+\Delta t)$ :

$$
\begin{align*}
& E\left(X_{1}(t+\Delta t) \mid X_{1}(t), X_{2}(t)\right)=X_{1}(t) e^{-c_{1} \Delta t}+X_{2}(t)\left(1-e^{-c_{2} \Delta t}\right)  \tag{4.33}\\
& E\left(X_{1}^{2}(t+\Delta t) \mid X_{1}(t), X_{2}(t)\right)=\begin{array}{l} 
\\
\\
\\
\\
\\
\\
\\
X_{2}(t) e^{-c_{1} \Delta t}\left(1-e^{-c_{1} \Delta t}\left(1-e^{-c_{2} \Delta t}\right)+\right. \\
\left.E^{2}(t+\Delta t) \mid X_{1}(t), X_{2}(t)\right)
\end{array}
\end{align*}
$$

After taking the expectations $E_{X_{1}(t), X_{2}(t)}(\cdot)$ of (4.33) and (4.34), we obtain the recursive expressions for the mean and the variance of the number of $S_{1}$ molecules

$$
\begin{align*}
\mu_{1}(t+\Delta t)= & \mu_{1}(t) e^{-c_{1} \Delta t}+\mu_{2}(t)\left(1-e^{-c_{2} \Delta t}\right)  \tag{4.35}\\
\sigma_{1}^{2}(t+\Delta t)= & E\left(X_{1}^{2}(t+\Delta t)\right)-\mu_{1}^{2}(t+\Delta t) \\
= & \mu_{1}(t) e^{-c_{1} \Delta t}\left(1-e^{-c_{1} \Delta t}\right)+\mu_{2}(t) e^{-c_{2} \Delta t}\left(1-e^{-c_{2} \Delta t}\right) \\
& +\sigma_{1}^{2}(t)\left(e^{-c_{1} \Delta t}\right)^{2}+\sigma_{2}^{2}(t)\left(1-e^{-c_{2} \Delta t}\right)^{2} \\
& -2 \sigma_{1}(t) \sigma_{2}(t) e^{-c_{1} \Delta t}\left(1-e^{-c_{2} \Delta t}\right) \tag{4.36}
\end{align*}
$$

where $\mu_{i}(t), i=1,2$ and $\sigma_{i}^{2}(t), i=1,2$ are the means and the variances of species $S_{1}$ and $S_{2}$, respectively.

### 4.3 General Expressions and Implementation

In order to implement the RM method, we choose a time step $\Delta t$ such that the probability of each reaction is less than a predefined value. That is

$$
\begin{equation*}
c_{m} \Delta t \leq \epsilon, \quad \forall m=1, \ldots M \tag{4.37}
\end{equation*}
$$

where $\epsilon$ is some specified value. Given the reaction stoichiometries, reaction constants, the initial number of molecules $\boldsymbol{X}(t=0)$ and the stop time $T_{s}$, the algorithm is implemented as follows:

## Initialization

- Select the time step as described above.
- Compute the matrix of transition probabilities $\boldsymbol{P}$ as given by the expressions (4.6a) and (4.6b).
- Set time $t=0$,
$\boldsymbol{\mu}(t=0)=\boldsymbol{X}(t=0)$, $\boldsymbol{C}(t=0)=\mathbf{0}$


## Recursion

Repeat the following steps until stop time $t=T_{s}$ is reached:

- Compute the matrix $\boldsymbol{B}(t)$ by employing the expressions (4.25a) and (4.25b).
- Compute $\boldsymbol{\mu}(t+\Delta t)$ by employing expression (4.17).
- Compute $\boldsymbol{C}(t+\Delta t)$ by employing expression (4.29).
- Advance the time to $t+\Delta t$.


### 4.4 Simulation Results

As described in Section 3.2, the SSA method generates realizations from the exact distribution. In this section, we compare results from the RM method with the SSA method
for accuracy and computation time. We track the means and the covariance matrix with time. In some cases, at a given time $t$, a normal distribution $\mathcal{N}\left(\mu_{i}(t), \sigma_{i}^{2}(t)\right)$ for species $S_{i}$, using the mean and the variance computed from the RM method, is compared with the distribution generated by the SSA method. The timing measurements are done on a machine called Fermat which consists of an Intel Pentium(R) Dual 2.4 GHz CPU.

### 4.4.1 Example - Three Species

Consider the reaction system

$$
\begin{align*}
& R_{1}: S_{1} \xrightarrow{c_{1}} S_{2}  \tag{4.38a}\\
& R_{2}: S_{2} \xrightarrow{c_{2}} S_{3} . \tag{4.38b}
\end{align*}
$$

We track all three species, and therefore $N=3$ and the number of reactions $M=2$. The state of the system is given by $\boldsymbol{X}(t)=\left(X_{1}(t) X_{2}(t) X_{3}(t)\right)^{T}$. The model is specified by the initial number of molecules $\boldsymbol{X}(t=0)=\left(\begin{array}{lll}10000 & 500 & 0\end{array}\right)^{T}$ and the reaction constants $c_{1}=0.1$ $\mathrm{sec}^{-1}$ and $c_{2}=1.0 \mathrm{sec}^{-1}$. The matrix of transition probabilities is defined by

$$
\boldsymbol{P}=\left(\begin{array}{ccc}
e^{-c_{1} \Delta t} & 1-e^{-c_{1} \Delta t} & 0  \tag{4.39}\\
0 & e^{-c_{2} \Delta t} & 1-e^{-c_{2} \Delta t} \\
0 & 0 & 1
\end{array}\right)
$$

$\boldsymbol{B}(t)$ is symmetric and its elements are given by

$$
\begin{aligned}
b_{11}(t) & =\mu_{1}(t) p_{11}\left(1-p_{11}\right) \\
b_{12}(t) & =\mu_{1}(t) p_{11} p_{12} \\
b_{13}(t) & =b_{31}(t)=0 \\
b_{22}(t) & =\mu_{1}(t) p_{12}\left(1-p_{12}\right)+\mu_{2}(t) p_{22}\left(1-p_{22}\right) \\
b_{23}(t) & =b_{32}(t)=\mu_{2}(t) p_{22} p_{23} \\
b_{33}(t) & =\mu_{2}(t) p_{23}\left(1-p_{23}\right) .
\end{aligned}
$$

By applying equations (4.17) and (4.29) repeatedly, we obtain the mean and the covariance of all the species from time $t=0 \mathrm{sec}$ to $t=10 \mathrm{sec}$. In Figure 4.1, we show the tracking results
of some of the first two moments of the species. In implementing the methods, $\Delta t=0.005$ sec for the RM method and the number of realizations is 5,000 for the SSA method. The average number of time steps with the SSA method is 12,721 . As seen from the figure, the estimated moments obtained by the RM and SSA methods are in good agreement.

### 4.4.2 Complex Example

We now present a complex example with the number of species $N=100$ and the number of reactions $M=200$. The reactions are listed in Tables 4.1 and 4.2, and the reaction parameters are listed in Table 4.3. The initial population $X_{1}(t=0)$ of the number of $S_{1}$ molecules is 2,000 . The initial populations of all the other species are zero. The time step $\Delta t$ in the RM method was selected to be $\Delta t=0.001 \mathrm{sec}$. The average number of time steps with the SSA method is 30,663 . We compared the results with the SSA for 500 realizations (Figures 4.2 and 4.3) and 2, 000 realizations (Figures 4.4 and 4.5). The SSA method produces much smoother results with 2,000 realizations and the results from the two methods compare well.

Next, we compare the computation time of the RM method with the SSA method with varying molecular populations. We vary the initial population $X_{1}(t=0)$ of the number of $S_{1}$ molecules from 5, 000 to 30,000 . The initial populations of all the other species are zero. As seen from (4.37), the time step $\Delta t$ of the RM method does not depend on the size of molecular population, and it is set to $\Delta t=0.001 \mathrm{sec}$ as indicated above.

For the SSA method, the simulation time depends on the molecular population, as the time to next reaction has an exponential distribution with an expected value $1 / a_{0}$ (expression (3.3)). Therefore, the number of time steps and hence the computation time increases with the molecular population. The number of realizations required to estimate the means and the variances also depend on the size of molecular populations. To achieve the same accuracy, we set the number of realizations equal to the molecular populations. Table 4.4 lists the computation times from both methods on the Fermat machine. With the SSA method, the CPU time is linear with the number of realizations. It was measured for a smaller number of realizations, and estimated for the number of realizations indicated in the table. The cumulative increase in the SSA CPU time is quadratic in the molecular populations. With the RM method the CPU time remains constant at 35 sec . We would like to remark that
the computation time employing the RM method will increase with the number of species $N$, as this requires multiplication of matrices of size $N$. Also for a given number of species, the computation time with the RM method remains constant as the number of reactions $M$ increases, whereas it increases with the SSA method.

| $S_{1} \longrightarrow S_{2}$ | $S_{2} \longrightarrow S_{3}$ | $S_{3} \longrightarrow S_{4}$ | $S_{4} \longrightarrow S_{5}$ | $S_{5} \longrightarrow S_{6}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $S_{6} \longrightarrow S_{7}$ | $S_{7} \longrightarrow S_{8}$ | $S_{8} \longrightarrow S_{9}$ | $S_{9} \longrightarrow S_{10}$ | $S_{10} \longrightarrow S_{11}$ |
| $S_{11} \longrightarrow S_{12}$ | $S_{12} \longrightarrow S_{13}$ | $S_{13} \longrightarrow S_{14}$ | $S_{14} \longrightarrow S_{15}$ | $S_{15} \longrightarrow S_{16}$ |
| $S_{16} \longrightarrow S_{17}$ | $S_{17} \longrightarrow S_{18}$ | $S_{18} \longrightarrow S_{19}$ | $S_{19} \longrightarrow S_{20}$ | $S_{20} \longrightarrow S_{21}$ |
| $S_{21} \longrightarrow S_{22}$ | $S_{22} \longrightarrow S_{23}$ | $S_{23} \longrightarrow S_{24}$ | $S_{24} \longrightarrow S_{25}$ | $S_{25} \longrightarrow S_{26}$ |
| $S_{26} \longrightarrow S_{27}$ | $S_{27} \longrightarrow S_{28}$ | $S_{28} \longrightarrow S_{29}$ | $S_{29} \longrightarrow S_{30}$ | $S_{30} \longrightarrow S_{31}$ |
| $S_{31} \longrightarrow S_{32}$ | $S_{32} \longrightarrow S_{33}$ | $S_{33} \longrightarrow S_{34}$ | $S_{34} \longrightarrow S_{35}$ | $S_{35} \longrightarrow S_{36}$ |
| $S_{36} \longrightarrow S_{37}$ | $S_{37} \longrightarrow S_{38}$ | $S_{38} \longrightarrow S_{39}$ | $S_{39} \longrightarrow S_{40}$ | $S_{40} \longrightarrow S_{41}$ |
| $S_{41} \longrightarrow S_{42}$ | $S_{42} \longrightarrow S_{43}$ | $S_{43} \longrightarrow S_{44}$ | $S_{44} \longrightarrow S_{45}$ | $S_{45} \longrightarrow S_{46}$ |
| $S_{46} \longrightarrow S_{47}$ | $S_{47} \longrightarrow S_{48}$ | $S_{48} \longrightarrow S_{49}$ | $S_{49} \longrightarrow S_{50}$ | $S_{50} \longrightarrow S_{51}$ |
| $S_{51} \longrightarrow S_{52}$ | $S_{52} \longrightarrow S_{53}$ | $S_{53} \longrightarrow S_{54}$ | $S_{54} \longrightarrow S_{55}$ | $S_{55} \longrightarrow S_{56}$ |
| $S_{56} \longrightarrow S_{57}$ | $S_{57} \longrightarrow S_{58}$ | $S_{58} \longrightarrow S_{59}$ | $S_{59} \longrightarrow S_{60}$ | $S_{60} \longrightarrow S_{61}$ |
| $S_{61} \longrightarrow S_{62}$ | $S_{62} \longrightarrow S_{63}$ | $S_{63} \longrightarrow S_{64}$ | $S_{64} \longrightarrow S_{65}$ | $S_{65} \longrightarrow S_{66}$ |
| $S_{66} \longrightarrow S_{67}$ | $S_{67} \longrightarrow S_{68}$ | $S_{68} \longrightarrow S_{69}$ | $S_{69} \longrightarrow S_{70}$ | $S_{70} \longrightarrow S_{71}$ |
| $S_{71} \longrightarrow S_{72}$ | $S_{72} \longrightarrow S_{73}$ | $S_{73} \longrightarrow S_{74}$ | $S_{74} \longrightarrow S_{75}$ | $S_{75} \longrightarrow S_{76}$ |
| $S_{76} \longrightarrow S_{77}$ | $S_{77} \longrightarrow S_{78}$ | $S_{78} \longrightarrow S_{79}$ | $S_{79} \longrightarrow S_{80}$ | $S_{80} \longrightarrow S_{81}$ |
| $S_{81} \longrightarrow S_{82}$ | $S_{82} \longrightarrow S_{83}$ | $S_{83} \longrightarrow S_{84}$ | $S_{84} \longrightarrow S_{85}$ | $S_{85} \longrightarrow S_{86}$ |
| $S_{86} \longrightarrow S_{87}$ | $S_{87} \longrightarrow S_{88}$ | $S_{88} \longrightarrow S_{89}$ | $S_{89} \longrightarrow S_{90}$ | $S_{90} \longrightarrow S_{91}$ |
| $S_{91} \longrightarrow S_{92}$ | $S_{92} \longrightarrow S_{93}$ | $S_{93} \longrightarrow S_{94}$ | $S_{94} \longrightarrow S_{95}$ | $S_{95} \longrightarrow S_{96}$ |
| $S_{96} \longrightarrow S_{97}$ | $S_{97} \longrightarrow S_{98}$ | $S_{98} \longrightarrow S_{99}$ | $S_{99} \longrightarrow S_{100}$ | $S_{100} \longrightarrow S_{1}$ |

Table 4.1: Complex Example, reactions $R_{1}$ through $R_{100}$, in left-to-right top-to-bottom order.

| $S_{1} \longrightarrow S_{11}$ | $S_{2} \longrightarrow S_{12}$ | $S_{3} \longrightarrow S_{13}$ | $S_{4} \longrightarrow S_{14}$ | $S_{5} \longrightarrow S_{15}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $S_{6} \longrightarrow S_{16}$ | $S_{7} \longrightarrow S_{17}$ | $S_{8} \longrightarrow S_{18}$ | $S_{9} \longrightarrow S_{19}$ | $S_{10} \longrightarrow S_{20}$ |
| $S_{11} \longrightarrow S_{21}$ | $S_{12} \longrightarrow S_{22}$ | $S_{13} \longrightarrow S_{23}$ | $S_{14} \longrightarrow S_{24}$ | $S_{15} \longrightarrow S_{25}$ |
| $S_{16} \longrightarrow S_{26}$ | $S_{17} \longrightarrow S_{27}$ | $S_{18} \longrightarrow S_{28}$ | $S_{19} \longrightarrow S_{29}$ | $S_{20} \longrightarrow S_{30}$ |
| $S_{21} \longrightarrow S_{31}$ | $S_{22} \longrightarrow S_{32}$ | $S_{23} \longrightarrow S_{33}$ | $S_{24} \longrightarrow S_{34}$ | $S_{25} \longrightarrow S_{35}$ |
| $S_{26} \longrightarrow S_{36}$ | $S_{27} \longrightarrow S_{37}$ | $S_{28} \longrightarrow S_{38}$ | $S_{29} \longrightarrow S_{39}$ | $S_{30} \longrightarrow S_{40}$ |
| $S_{31} \longrightarrow S_{41}$ | $S_{32} \longrightarrow S_{42}$ | $S_{33} \longrightarrow S_{43}$ | $S_{34} \longrightarrow S_{44}$ | $S_{35} \longrightarrow S_{45}$ |
| $S_{36} \longrightarrow S_{46}$ | $S_{37} \longrightarrow S_{47}$ | $S_{38} \longrightarrow S_{48}$ | $S_{39} \longrightarrow S_{49}$ | $S_{40} \longrightarrow S_{50}$ |
| $S_{41} \longrightarrow S_{51}$ | $S_{42} \longrightarrow S_{52}$ | $S_{43} \longrightarrow S_{53}$ | $S_{44} \longrightarrow S_{54}$ | $S_{45} \longrightarrow S_{55}$ |
| $S_{46} \longrightarrow S_{56}$ | $S_{47} \longrightarrow S_{57}$ | $S_{48} \longrightarrow S_{58}$ | $S_{49} \longrightarrow S_{59}$ | $S_{50} \longrightarrow S_{60}$ |
| $S_{51} \longrightarrow S_{61}$ | $S_{52} \longrightarrow S_{62}$ | $S_{53} \longrightarrow S_{63}$ | $S_{54} \longrightarrow S_{64}$ | $S_{55} \longrightarrow S_{65}$ |
| $S_{56} \longrightarrow S_{66}$ | $S_{57} \longrightarrow S_{67}$ | $S_{58} \longrightarrow S_{68}$ | $S_{59} \longrightarrow S_{69}$ | $S_{60} \longrightarrow S_{70}$ |
| $S_{61} \longrightarrow S_{71}$ | $S_{62} \longrightarrow S_{72}$ | $S_{63} \longrightarrow S_{73}$ | $S_{64} \longrightarrow S_{74}$ | $S_{65} \longrightarrow S_{75}$ |
| $S_{66} \longrightarrow S_{76}$ | $S_{67} \longrightarrow S_{77}$ | $S_{68} \longrightarrow S_{78}$ | $S_{69} \longrightarrow S_{79}$ | $S_{70} \longrightarrow S_{80}$ |
| $S_{71} \longrightarrow S_{81}$ | $S_{72} \longrightarrow S_{82}$ | $S_{73} \longrightarrow S_{83}$ | $S_{74} \longrightarrow S_{84}$ | $S_{75} \longrightarrow S_{85}$ |
| $S_{76} \longrightarrow S_{86}$ | $S_{77} \longrightarrow S_{87}$ | $S_{78} \longrightarrow S_{88}$ | $S_{79} \longrightarrow S_{89}$ | $S_{80} \longrightarrow S_{90}$ |
| $S_{81} \longrightarrow S_{91}$ | $S_{82} \longrightarrow S_{92}$ | $S_{83} \longrightarrow S_{93}$ | $S_{84} \longrightarrow S_{94}$ | $S_{85} \longrightarrow S_{95}$ |
| $S_{86} \longrightarrow S_{96}$ | $S_{87} \longrightarrow S_{97}$ | $S_{88} \longrightarrow S_{98}$ | $S_{89} \longrightarrow S_{99}$ | $S_{90} \longrightarrow S_{100}$ |
| $S_{91} \longrightarrow S_{1}$ | $S_{92} \longrightarrow S_{1}$ | $S_{93} \longrightarrow S_{3}$ | $S_{94} \longrightarrow S_{4}$ | $S_{95} \longrightarrow S_{5}$ |
| $S_{96} \longrightarrow S_{6}$ | $S_{97} \longrightarrow S_{7}$ | $S_{98} \longrightarrow S_{8}$ | $S_{99} \longrightarrow S_{9}$ | $S_{100} \longrightarrow S_{10}$ |

Table 4.2: Complex Example, reactions $R_{101}$ through $R_{200}$, in left-to-right top-to-bottom order.

| 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 |
| 0.03 | 0.03 | 0.03 | 0.03 | 0.03 | 0.03 | 0.03 | 0.03 | 0.03 | 0.03 |
| 0.04 | 0.04 | 0.04 | 0.04 | 0.04 | 0.04 | 0.04 | 0.04 | 0.04 | 0.04 |
| 0.05 | 0.05 | 0.05 | 0.05 | 0.05 | 0.05 | 0.05 | 0.05 | 0.05 | 0.05 |
| 0.06 | 0.06 | 0.06 | 0.06 | 0.06 | 0.06 | 0.06 | 0.06 | 0.06 | 0.06 |
| 0.07 | 0.07 | 0.07 | 0.07 | 0.07 | 0.07 | 0.07 | 0.07 | 0.07 | 0.07 |
| 0.08 | 0.08 | 0.08 | 0.08 | 0.08 | 0.08 | 0.08 | 0.08 | 0.08 | 0.08 |
| 0.09 | 0.09 | 0.09 | 0.09 | 0.09 | 0.09 | 0.09 | 0.09 | 0.09 | 0.09 |
| 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 |
| 1.01 | 1.01 | 1.01 | 1.01 | 1.01 | 1.01 | 1.01 | 1.01 | 1.01 | 1.01 |
| 1.02 | 1.02 | 1.02 | 1.02 | 1.02 | 1.02 | 1.02 | 1.02 | 1.02 | 1.02 |
| 1.03 | 1.03 | 1.03 | 1.03 | 1.03 | 1.03 | 1.03 | 1.03 | 1.03 | 1.03 |
| 1.04 | 1.04 | 1.04 | 1.04 | 1.04 | 1.04 | 1.04 | 1.04 | 1.04 | 1.04 |
| 1.05 | 1.05 | 1.05 | 1.05 | 1.05 | 1.05 | 1.05 | 1.05 | 1.05 | 1.05 |
| 1.06 | 1.06 | 1.06 | 1.06 | 1.06 | 1.06 | 1.06 | 1.06 | 1.06 | 1.06 |
| 1.07 | 1.07 | 1.07 | 1.07 | 1.07 | 1.07 | 1.07 | 1.07 | 1.07 | 1.07 |
| 1.08 | 1.08 | 1.08 | 1.08 | 1.08 | 1.08 | 1.08 | 1.08 | 1.08 | 1.08 |
| 1.09 | 1.09 | 1.09 | 1.09 | 1.09 | 1.09 | 1.09 | 1.09 | 1.09 | 1.09 |
| 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
|  |  |  |  |  |  |  |  |  |  |

Table 4.3: Complex Example, reaction constants for reactions $R_{1}$ through $R_{200}$, in units of $\mathrm{sec}^{-1}$, in left-to-right top-to-bottom order.

| Initial number of <br> $S_{1}$ molecules | Number of realizations <br> in SSA | CPU time |  |
| :---: | :---: | :---: | :---: |
|  |  | SSA |  |
| 5,000 | 5,000 | 35 sec | 2.35 days |
| 10,000 | 10,000 | 35 sec | 8.9 days |
| 15,000 | 15,000 | 35 sec | 19.2 days |
| 20,000 | 20,000 | 35 sec | 36.6 days |
| 25,000 | 25,000 | 35 sec | 61.1 days |
| 30,000 | 30,000 | 35 sec | 85.1 days |

Table 4.4: Computation times of the RM and the SSA methods, versus the molecular populations, for the Complex Example given in Section 4.4.2.


Figure 4.1: Comparison of the results from the RM and the SSA methods for the Three Species example given in Section 4.4.1. The curves represent the computed means $\mu_{2}(t)$ (top plot), the variances $\sigma_{2}^{2}(t)$ (second plot), and the covariances $\sigma_{21}(t)$ (third plot) and $\sigma_{23}(t)$ (bottom plot). $\Delta t$ is .005 sec with the RM method and the number of realizations in the SSA method is 5,000 .


Figure 4.2: Comparison of the results from the RM and the SSA methods for the Complex Example given in Section 4.4.2. The curves represent the computed means $\mu_{11}(t)$ (top plot), the variances $\sigma_{11}^{2}(t)$ (second plot), and the covariances $\sigma_{11,10}(t)$ (third plot) and $\sigma_{11,12}(t)$ (bottom plot). $\Delta t$ is .0005 sec with the RM method and the number of realizations in the SSA method is 1,000 .


Figure 4.3: The normal distribution generated from the RM method (solid line) is compared with the distribution generated by the SSA method ("noisy" plot), at time $t=10 \mathrm{sec}$, for the Complex Example given in Section 4.4.2.


Figure 4.4: Same example and parameters as in Figure 4.2, except that the number of realizations in the SSA method is 2,000 .


Figure 4.5: Same example and parameters as in Figure 4.3, except that the number of realizations in the SSA method is 2,000 .

## Chapter 5

## Proposed Method - Second Order Reactions

### 5.1 Introduction

In Chapter 4, we obtained recursive expressions for propagating the first two moments of the joint distribution of the molecular species. Our results were applied to a complex system consisting of a hundred species and two hundred reactions. We now want to extend our method to the more difficult problem of second order reactions. Some issues with the second order reactions are:

- In first order reactions, the reaction rates are proportional to the number of molecules of the reactant. In second order reactions, the reaction rates are proportional to the product of the number of molecules of the reactants. That is, the second order reactions are nonlinear in propensities.
- The evaluation of the covariance matrix requires computing expectations of third and higher order cross terms.
- For the second order reaction $S_{i}+S_{j} \longrightarrow S_{k}$, the expressions (2.29) for $E\left(X_{1}(t)\right)$ and (2.30) containing $E\left(X_{1}^{2}(t)\right)$ do not have a simple form as for the first order reaction $S_{i} \longrightarrow S_{j}$.


### 5.2 The Methodology

Once again, we use the conditional expectation relations as in (4.14) and (4.15)

$$
\begin{align*}
& \boldsymbol{\mu}(t+\Delta t) \equiv E_{\boldsymbol{X}(t+\Delta t)}(\boldsymbol{X}(t+\Delta t)) \\
& =E_{\boldsymbol{X}(t)}\left(E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t))\right)  \tag{5.1}\\
& E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}\left(\boldsymbol{X}(t+\Delta t) \boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \\
& =\operatorname{cov}\left(\boldsymbol{X}(t+\Delta t), \boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \\
& \quad+E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)) E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}\left(\boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) . \tag{5.2}
\end{align*}
$$

For the second order reaction

$$
\begin{equation*}
R_{m}: S_{i}+S_{j} \xrightarrow{c_{m}} S_{k}, \tag{5.3}
\end{equation*}
$$

we do not have a simple expression for the transition probabilities of the above reaction. We use the approximation that in the time interval $\Delta t$, the probability that a molecule of species $S_{i}$ (or $S_{j}$ ) changes into a molecule of species $S_{k}$ is $c_{m} \Delta t$, and the probability that a molecule of species $S_{i}$ ( or $S_{j}$ ) remains a molecule of species $S_{i}$ (or $S_{j}$ ) is $1-c_{m} \Delta t$. We also have reactions of the form

$$
\begin{equation*}
2 S_{1} \longrightarrow S_{2}, \tag{5.4}
\end{equation*}
$$

where the stoichiometry of species $S_{1}$ is -2 and

$$
\begin{equation*}
S_{1}+S_{2} \longrightarrow 2 S_{3}, \tag{5.5}
\end{equation*}
$$

where the stoichiometry of species $S_{3}$ is 2 . Therefore, in the following derivations we take into account $v_{m i}$, the stoichiometry of species $S_{i}$ in reaction $R_{m}$. We define a matrix $\boldsymbol{W}$ of size $M \times N$ with elements

$$
\begin{equation*}
w_{m i}=v_{m i}\left(c_{m} \Delta t\right), \quad m=1,2 \ldots M, \quad i=1,2 \ldots N . \tag{5.6}
\end{equation*}
$$

We further define a vector $\boldsymbol{H}(t)$ of size $M$ such that its $m$ th element $h_{m}(t)$ is the number
of distinct molecular combinations for reaction $R_{m}$ as described in Section 1.4.2. We obtain an expression for $E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t))$ by summing up the contribution for each species from all the reactions. This gives

$$
\begin{equation*}
E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t))=\boldsymbol{X}(t)+\boldsymbol{W}^{T} \boldsymbol{H}(t) . \tag{5.7}
\end{equation*}
$$

Taking expectation $E_{\boldsymbol{X}(t)}(\cdot)$ of the above expression and setting $\boldsymbol{G}(t) \equiv E_{\boldsymbol{X}(t)}(\boldsymbol{H}(t))$, we obtain

$$
\begin{equation*}
\boldsymbol{\mu}(t+\Delta t)=\boldsymbol{\mu}(t)+\boldsymbol{W}^{T} \boldsymbol{G}(t) \tag{5.8}
\end{equation*}
$$

We next evaluate the recursive expression for the covariance matrix. Utilizing (5.7) yields the second term in (5.2)

$$
\begin{align*}
& E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}(\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)) E_{\boldsymbol{X}(t+\Delta t) \mid \boldsymbol{X}(t)}\left(\boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \\
& =\boldsymbol{X}(t) \boldsymbol{X}^{T}(t)+\boldsymbol{X}(t) \boldsymbol{H}^{T}(t) \boldsymbol{W} \\
& \quad+\boldsymbol{W}^{T} \boldsymbol{H}(t) \boldsymbol{X}^{T}(t)+\boldsymbol{W}^{T} \boldsymbol{H}(t) \boldsymbol{H}^{T}(t) \boldsymbol{W} . \tag{5.9}
\end{align*}
$$

To evaluate the first term in (5.2) we set

$$
\begin{equation*}
\boldsymbol{R}(t) \equiv \operatorname{cov}\left(\boldsymbol{X}(t+\Delta t), \boldsymbol{X}^{T}(t+\Delta t) \mid \boldsymbol{X}(t)\right) \tag{5.10}
\end{equation*}
$$

where $\boldsymbol{R}(t)$ is a matrix of size $N$ and its elements are obtained using the multinomial distribution approximation, as in the case of first order reactions. That is, in the time interval $(t, t+\Delta t)$, the reactant species in the reaction channel $R_{m}$ either remain as reactants or change into its product species, as specified by the stoichiometries. Further, the reaction propensities are proportional to $h_{m}(t)$ for second order reactions, and therefore we replace $K_{m}$ by $h_{m}(t)$ in expressions (4.22) and (4.23). This yields

$$
\begin{equation*}
r_{i i}(t)=\sum_{m=1}^{M} v_{m i}^{2} h_{m}(t) p_{m i}\left(1-p_{m i}\right) \tag{5.11a}
\end{equation*}
$$

$$
\begin{equation*}
r_{i j}(t)=\sum_{m=1}^{M} v_{m i} v_{m j} h_{m}(t) p_{m i} p_{m j} \quad i \neq j \tag{5.11b}
\end{equation*}
$$

where $p_{m i}=1-c_{m} \Delta t$, if species $S_{i}$ is a reactant and $p_{m i}=c_{m} \Delta t$ if $S_{i}$ is a product in reaction $R_{m}$. Marginalizing the above expression gives the elements of $\boldsymbol{B}(t)=E_{\boldsymbol{X}(t)}(\boldsymbol{R}(t))$

$$
\begin{align*}
& b_{i i}(t)=\sum_{m=1}^{M} v_{m i}^{2} g_{m}(t) p_{m i}\left(1-p_{m i}\right)  \tag{5.12a}\\
& b_{i j}(t)=\sum_{m=1}^{M} v_{m i} v_{m j} g_{m}(t) p_{m i} p_{m j} \tag{5.12b}
\end{align*}
$$

where $g_{m}(t)$ are the elements of $\boldsymbol{G}(t)$.
Substituting (5.9) and (5.10) in (5.2), and taking its expectation $E_{\boldsymbol{X}(t)}(\cdot)$ gives the second moments

$$
\begin{align*}
E(\boldsymbol{X}(t+\Delta t) & \left.\boldsymbol{X}^{T}(t+\Delta t)\right) \\
= & E_{\boldsymbol{X}(t)}\left(\boldsymbol{X}(t) \boldsymbol{X}^{T}(t)\right)+\boldsymbol{B}(t)+E_{\boldsymbol{X}(t)}\left(\boldsymbol{X}(t) \boldsymbol{H}^{T}(t)\right) \boldsymbol{W} \\
& +\boldsymbol{W}^{T} E_{\boldsymbol{X}(t)}\left(\boldsymbol{H}(t) \boldsymbol{X}^{T}(t)\right)+\boldsymbol{W}^{T} E_{\boldsymbol{X}(t)}\left(\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)\right) \boldsymbol{W} \tag{5.13}
\end{align*}
$$

From (5.8) and (5.13), we obtain the following recursive expression for the covariance matrix

$$
\begin{align*}
\boldsymbol{C}(t+\Delta t)= & \boldsymbol{C}(t)+\boldsymbol{B}(t)+\left(E_{\boldsymbol{X}(t)}\left(\boldsymbol{X}(t) \boldsymbol{H}^{T}(t)\right)-\boldsymbol{\mu}(t) \boldsymbol{G}^{T}(t)\right) \boldsymbol{W} \\
& +\boldsymbol{W}^{T}\left(E_{\boldsymbol{X}(t)}\left(\boldsymbol{H}(t) \boldsymbol{X}^{T}(t)\right)-\boldsymbol{G}(t) \boldsymbol{\mu}^{T}(t)\right) \\
& +\boldsymbol{W}^{T}\left(E_{\boldsymbol{X}(t)}\left(\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)\right)-\boldsymbol{G}(t) \boldsymbol{G}^{T}(t)\right) \boldsymbol{W} \tag{5.14}
\end{align*}
$$

where the fourth term is the transpose of the third term.

### 5.2.1 Example - Simple Reversible Reaction

Consider the reactions

$$
\begin{align*}
& R_{1}: S_{1}+S_{2} \xrightarrow{c_{1}} S_{3}  \tag{5.15a}\\
& R_{2}: S_{3} \xrightarrow{c_{2}} S_{1}+S_{2} . \tag{5.15b}
\end{align*}
$$

If we know the initial number of molecules of all the species, then it is sufficient to track only one species. We will track species, $S_{1}$, and therefore we have $N=1, M=2$ and $\boldsymbol{X}(t)=X_{1}(t)$. The stoichiometries of species $S_{1}$ in reactions $R_{1}$ and $R_{2}$ are $\nu_{11}=-1$ and $\nu_{21}=1$, respectively. The matrices $\boldsymbol{H}(t)$ and $\boldsymbol{W}$ are given by

$$
\begin{equation*}
\boldsymbol{H}(t)=\binom{X_{1}(t) X_{2}(t)}{X_{3}(t)}, \quad \boldsymbol{W}=\binom{-c_{1} \Delta t}{c_{2} \Delta t} \tag{5.16}
\end{equation*}
$$

The matrix $\boldsymbol{R}$ is of size one, and its element is given by summing up the contributions from reactions $R_{1}$ and $R_{2}$

$$
\begin{equation*}
r_{11}=X_{1}(t) X_{2}(t) c_{1} \Delta t\left(1-c_{1} \Delta t\right)+X_{3}(t) c_{2} \Delta t\left(1-c_{2} \Delta t\right) . \tag{5.17}
\end{equation*}
$$

Utilizing (5.8) and (5.14) and expanding gives us the following expressions for the propagation of the mean and the variance of $X_{1}(t)$ :

$$
\begin{align*}
\mu_{1}(t+\Delta t)= & \mu_{1}(t)-E\left(X_{1}(t) X_{2}(t)\right) c_{1} \Delta t+E\left(X_{3}(t)\right) c_{2} \Delta t  \tag{5.18}\\
\sigma_{1}^{2}(t+\Delta t)= & \sigma_{1}^{2}(t)+E\left(X_{1}(t) X_{2}(t)\right) c_{1} \Delta t\left(1-c_{1} \Delta t\right) \\
& +E\left(X_{3}(t)\right) c_{2} \Delta t\left(1-c_{2} \Delta t\right) \\
& +\left(c_{1} \Delta t\right)^{2}\left(E\left(X_{1}^{2}(t) X_{2}^{2}(t)\right)-\left(E\left(X_{1}(t) X_{2}(t)\right)\right)^{2}\right) \\
& +\left(c_{2} \Delta t\right)^{2}\left(E\left(X_{3}^{2}(t)\right)-\left(E\left(X_{3}(t)\right)\right)^{2}\right) \\
& -\left(2 c_{1} \Delta t\right)\left(E\left(X_{1}^{2}(t) X_{2}(t)\right)-E\left(X_{1}(t)\right) E\left(X_{1}(t) X_{2}(t)\right)\right) \\
& +\left(2 c_{2} \Delta t\right)\left(E\left(X_{1}(t) X_{3}(t)\right)-E\left(X_{1}(t)\right) E\left(X_{3}(t)\right)\right) \\
& -\left(2 c_{1} \Delta t c_{2} \Delta t\right)\left(E\left(X_{1}(t) X_{2}(t) X_{3}(t)\right)-E\left(X_{1}(t) X_{2}(t)\right) E\left(X_{3}(t)\right)\right) \tag{5.19}
\end{align*}
$$

In the above expressions $\mu_{i}(t)=E\left(X_{i}(t)\right), i=1,2,3$. In addition, we have the constraints $X_{1}(t)-X_{2}(t)=\beta_{12}, X_{1}(t)-X_{3}(t)=\beta_{13}$, on the molecular populations. The evaluation of higher order terms of the form $E\left(X_{1}^{2}(t) X_{2}(t)\right)$ will be discussed in Section 5.3.

### 5.2.2 Example - Competing Reactions

For the competing reactions

$$
\begin{align*}
& R_{1}: S_{1}+S_{2} \xrightarrow{c_{1}} S_{3}  \tag{5.20a}\\
& R_{2}: S_{1}+S_{4} \xrightarrow{c_{2}} S_{5} \tag{5.20b}
\end{align*}
$$

we track three species $S_{1}, S_{2}$ and $S_{4}$. The number of tracked species $N=3$, the number of reactions $M=2$ and $\boldsymbol{X}(t)=\left(\begin{array}{lll}X_{1}(t) & X_{2}(t) & X_{4}(t)\end{array}\right)^{T}$. The stoichiometries vectors are $\boldsymbol{\nu}_{1}=\left(\begin{array}{ll}-1-1 & 0\end{array}\right)$ and $\boldsymbol{\nu}_{2}=\left(\begin{array}{ll}-1 & 0-1\end{array}\right)$. The matrices $\boldsymbol{H}(t)$ and $\boldsymbol{W}$ are given by

$$
\boldsymbol{H}(t)=\binom{X_{1}(t) X_{2}(t)}{X_{1}(t) X_{4}(t)}, \quad \boldsymbol{W}=\left(\begin{array}{ccc}
-c_{1} \Delta t & -c_{1} \Delta t & 0  \tag{5.21}\\
-c_{2} \Delta t & 0 & -c_{2} \Delta t
\end{array}\right)
$$

The matrix

$$
\boldsymbol{R}(t)=\left(\begin{array}{ccc}
r_{11}(t) & r_{12}(t) & r_{14}(t)  \tag{5.22}\\
r_{21}(t) & r_{22}(t) & r_{24}(t) \\
r_{41}(t) & r_{42}(t) & r_{44}(t)
\end{array}\right)
$$

is symmetric and its elements are given by

$$
\begin{aligned}
& r_{11}(t)=X_{1}(t) X_{2}(t) c_{1} \Delta t\left(1-c_{1} \Delta t\right)+X_{1}(t) X_{4}(t) c_{2} \Delta t\left(1-c_{2} \Delta t\right) \\
& r_{12}(t)=X_{1}(t) X_{2}(t) c_{1} \Delta t\left(1-c_{1} \Delta t\right) \\
& r_{14}(t)=X_{1}(t) X_{4}(t) c_{2} \Delta t\left(1-c_{2} \Delta t\right) \\
& r_{22}(t)=X_{1}(t) X_{2}(t) c_{1} \Delta t\left(1-c_{1} \Delta t\right) \\
& r_{24}(t)=0 \\
& r_{44}(t)=X_{1}(t) X_{4}(t) c_{2} \Delta t\left(1-c_{2} \Delta t\right)
\end{aligned}
$$

This example will be continued in Section 5.5.1.

### 5.3 Higher Order Joint Moments

In order to evaluate the expression (5.14), we need to compute the expectations of $\boldsymbol{X}(t) \boldsymbol{H}^{T}(t)$ and $\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)$. These matrices consist of third and higher order terms of the form $X_{1}(t) X_{2}(t) X_{3}(t), X_{1}^{2}(t) X_{2}^{2}(t)$ and so forth. In general, the expressions for moments of order up to $Q$ consist of moments of order greater than $Q$. This is the moment closure problem and a suitable approximation has to be constructed. We will utilize the multivariate normal distribution approximation for each time step $\Delta t$. Gómez-Uribe et al. [32], Goutsias [35] and Lee et al. [47] have implemented moment closure by setting higher order moment terms to zero. It is true that for multivariate normal distribution, the central moments of odd order are zero, and third and higher order cumulants are zero, but the higher order moments contained in the expressions for computations of second moments (covariance matrix in (5.14)) are not necessarily third order central moments or third order cumulants. Singh et al. implemented a moment closure technique and obtained expressions for higher order moments that are consistent with lognormal distribution [68], [69]. We will compare the normal and lognormal moment closure expressions for an example presented in [38].

We now derive the moments of multivariate normal distribution, using the moment generating function [36], [74]. The multivariate normal PDF is defined by

$$
\begin{equation*}
f_{\boldsymbol{X}(t)}(\boldsymbol{x}(t))=\frac{1}{(2 \pi)^{N / 2}|\boldsymbol{C}(t)|^{1 / 2}} \exp \left\{-\frac{1}{2}(\boldsymbol{x}(t)-\boldsymbol{\mu}(t))^{T} \boldsymbol{C}^{-1}(t)(\boldsymbol{x}(t)-\boldsymbol{\mu}(t))\right\} \tag{5.23}
\end{equation*}
$$

Its moment generating function is given by

$$
\begin{equation*}
\mathcal{M}_{\boldsymbol{X}(t)}(\boldsymbol{\omega})=\exp \left\{\boldsymbol{\mu}^{T}(t) \boldsymbol{\omega}+\frac{1}{2} \boldsymbol{\omega}^{T} \boldsymbol{C}(t) \boldsymbol{\omega}\right\} \tag{5.24}
\end{equation*}
$$

where $\mathcal{M}_{\boldsymbol{X}(t)}(\boldsymbol{\omega})$ is defined by

$$
\begin{equation*}
\mathcal{M}_{\boldsymbol{X}(t)}(\boldsymbol{\omega})=E\left(\exp \left\{\boldsymbol{\omega}^{T} \boldsymbol{X}(t)\right\}\right) \tag{5.25}
\end{equation*}
$$

and $\boldsymbol{\omega}=\left(\begin{array}{llll}\omega_{1} & \omega_{2} & \ldots & \omega_{N}\end{array}\right)$ is a vector of size $N$. Recall from (4.1) that the joint non-central moments are defined by

$$
\begin{equation*}
\gamma^{\boldsymbol{q}}(t)=E\left(\boldsymbol{X}^{\boldsymbol{q}}(t)\right) \tag{5.26}
\end{equation*}
$$

| $E\left(X_{i}^{3}(t)\right)$ | $\mu_{i}^{3}(t)+3 \mu_{i}(t) \sigma_{i}^{2}(t)$ |
| :--- | :--- |
| $E\left(X_{i}^{4}(t)\right)$ | $\mu_{i}^{4}(t)+6 \mu_{i}^{2}(t) \sigma_{i}^{2}(t)+3 \sigma_{i}^{4}(t)$ |
| $E\left(X_{i}^{2}(t) X_{j}(t)\right)$ | $2 \mu_{i}(t) \sigma_{i j}(t)+\left(\mu_{i}^{2}(t)+\sigma_{i}^{2}(t)\right) \mu_{j}(t)$ |
| $E\left(X_{i}(t) X_{j}(t) X_{k}(t)\right)$ | $\mu_{i}(t) \mu_{j}(t) \mu_{k}(t)+\sigma_{j k}(t) \mu_{i}(t)+\sigma_{i k}(t) \mu_{j}(t)+\sigma_{i j}(t) \mu_{k}(t)$ |
| $E\left(X_{i}^{2}(t) X_{j}^{2}(t)\right)$ | $4 \mu_{i}(t) \mu_{j}(t) \sigma_{i j}(t)+\mu_{i}^{2}(t) \mu_{j}^{2}(t)+\mu_{i}^{2}(t) \sigma_{j}^{2}(t)$ |
|  | $+\mu_{j}^{2}(t) \sigma_{i}^{2}(t)+2 \sigma_{i j}^{2}(t)+\sigma_{i}^{2}(t) \sigma_{j}^{2}(t)$ |
| $E\left(X_{i}^{2}(t) X_{j}(t) X_{k}(t)\right)$ | $\mu_{i}^{2}(t) \mu_{j}(t) \mu_{k}(t)+\mu_{i}^{2}(t) \sigma_{j k}(t)+2 \mu_{i}(t) \mu_{j}(t) \sigma_{i k}(t)$ |
|  | $+2 \mu_{i}(t) \mu_{k}(t) \sigma_{i j}(t)+\sigma_{i}^{2}(t) \mu_{j}(t) \mu_{k}(t)+\sigma_{j k}(t) \sigma_{i}^{2}(t)+2 \sigma_{i j}(t) \sigma_{i k}(t)$ |

Table 5.1: Higher order expectations expressed in terms of the first and second moments for the multivariate normal distribution.

These moments are computed by utilizing (5.24) according to:

$$
\begin{equation*}
\gamma^{\boldsymbol{q}}(t)=\left.\frac{\partial^{q_{1}+q_{2} \ldots q_{N}}}{\partial \omega_{1}^{q_{1}} \partial \omega_{2}^{q_{2}} \ldots \partial \omega_{N}^{q_{N}}} \mathcal{M}_{\boldsymbol{X}(t)}(\boldsymbol{\omega})\right|_{\boldsymbol{\omega}=\mathbf{0}} \tag{5.27}
\end{equation*}
$$

The joint moments required for the computation of expectations in $\boldsymbol{X}(t) \boldsymbol{H}^{T}(t)$ and $\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)$ are listed in Table 5.1.

### 5.4 Algorithm Implementation

We choose a time step $\Delta t$ such that the average number of reactions in each reaction channel is less than a specified value. For the three types of reactions

$$
\begin{align*}
& R_{1}: S_{1} \longrightarrow S_{2}  \tag{5.28a}\\
& R_{2}: 2 S_{1} \longrightarrow S_{2}  \tag{5.28b}\\
& R_{3}: S_{1}+S_{2} \longrightarrow S_{3} \tag{5.28c}
\end{align*}
$$

choose a time step $\Delta t$ such that it satisfies

$$
\begin{align*}
c_{1} \Delta t & \leq \epsilon  \tag{5.29a}\\
x_{1}(t) c_{2} \Delta t & \leq \epsilon  \tag{5.29b}\\
K_{m}(t) c_{3} \Delta t & \leq \epsilon, \quad K_{m}(t)=\max \left(x_{1}(t), x_{2}(t)\right), \tag{5.29c}
\end{align*}
$$

where $\epsilon$ is some specified value. Note that the time step maybe updated once every few recursions. For second order reactions, the time step depends on the molecular populations.

Given the reaction stoichiometries, reaction constants, the initial number of molecules $\boldsymbol{X}(t=0)$ and the stop time $T_{s}$, the algorithm is implemented as follows:

## Initialization

- Define the matrices $\boldsymbol{X}(t)$ and $\boldsymbol{H}(t)$, for time $t$.
- Set time $t=0$,
$\boldsymbol{\mu}(t=0)=\boldsymbol{X}(t=0)$, $\boldsymbol{C}(t=0)=\mathbf{0}$.


## Recursion

Repeat the following steps until stop time $t=T_{s}$ is reached:

- Update the step size $\Delta t$ as described above.
- Compute the elements of the matrix $\boldsymbol{W}$ as given by (5.6).
- Compute the elements of the matrix $\boldsymbol{R}(\boldsymbol{t})$ as given by (5.11a) and (5.11b).
- Evaluate the expectations $E_{\boldsymbol{X}(t)}(\boldsymbol{H}(t))$ and $E_{\boldsymbol{X}(t)}(\boldsymbol{R}(t))$.
- Evaluate the expectations $E_{\boldsymbol{X}(t)}\left(\boldsymbol{X}(t) \boldsymbol{H}^{T}(t)\right)$ and $E_{\boldsymbol{X}(t)}\left(\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)\right)$ from Table 5.1.
- Compute $\boldsymbol{\mu}(t+\Delta t)$ by employing expression (5.8).
- Compute $\boldsymbol{C}(t+\Delta t)$ by employing expression (5.14).
- Advance the time to $t+\Delta t$.


### 5.5 Simulation Results

As with the first order system, we compare the simulation results from the second order expressions with the SSA, for accuracy and computation time. We track the means and the covariance matrix with time. We also compare plots of normal distribution $\mathcal{N}\left(\mu_{i}(t), \sigma_{i}^{2}(t)\right)$, for species $S_{i}$ at time $t$, using the mean and the variance computed from the RM method, with the distribution generated by the SSA method.

### 5.5.1 Example - Competing Reactions (continued)

We continue the example from Section 5.2.2. We have

$$
\boldsymbol{\mu}(t)=\left(\begin{array}{l}
\mu_{1}(t)  \tag{5.30}\\
\mu_{2}(t) \\
\mu_{4}(t)
\end{array}\right) \quad \boldsymbol{C}(t)=\left(\begin{array}{ccc}
\sigma_{1}^{2}(t) & \sigma_{12}(t) & \sigma_{14}(t) \\
\sigma_{21}(t) & \sigma_{2}^{2}(t) & \sigma_{24}(t) \\
\sigma_{41}(t) & \sigma_{42}(t) & \sigma_{4}^{2}(t)
\end{array}\right) .
$$

In order to compute $\boldsymbol{\mu}(t+\Delta t)$ and $\boldsymbol{C}(t+\Delta t)$, we need to evaluate the expectations $\boldsymbol{G}(t) \equiv E_{\boldsymbol{X}(t)}(\boldsymbol{H}(t))$ and $\boldsymbol{B}(t)=E_{\boldsymbol{X}(t)}(\boldsymbol{R}(t))$. In addition, we need the expectations of $\boldsymbol{H}(t) \boldsymbol{X}^{T}(t), \boldsymbol{X}(t) \boldsymbol{H}^{T}(t)$ and $\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)$. The expressions for $\boldsymbol{X}(t) \boldsymbol{H}^{T}(t)$ and $\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)$ are

$$
\begin{gather*}
\boldsymbol{X}(t) \boldsymbol{H}^{T}(t)=\left(\begin{array}{cc}
X_{1}^{2}(t) X_{2}(t) & X_{1}^{2}(t) X_{4}(t) \\
X_{1}(t) X_{2}^{2}(t) & X_{1}(t) X_{2}(t) X_{4}(t) \\
X_{1}(t) X_{2}(t) X_{4}(t) & X_{1}(t) X_{4}^{2}(t)
\end{array}\right)  \tag{5.31}\\
\boldsymbol{H}(t) \boldsymbol{H}^{T}(t)=\left(\begin{array}{cc}
X_{1}^{2}(t) X_{2}^{2}(t) & X_{1}^{2}(t) X_{2}(t) X_{4}(t) \\
X_{1}^{2}(t) X_{2}(t) X_{4}(t) & X_{1}^{2}(t) X_{4}^{2}(t)
\end{array}\right) . \tag{5.32}
\end{gather*}
$$

The required expectations are evaluated using joint moments of the multivariate normal distribution as described in Section 5.3. Some of these terms are

$$
\begin{align*}
E\left(X_{1}(t) X_{2}(t)\right)= & \sigma_{12}(t)+\mu_{1}(t) \mu_{2}(t)  \tag{5.33}\\
E\left(X_{1}^{2}(t) X_{2}(t)\right)= & 2 \mu_{1}(t) \sigma_{12}(t)+\left(\mu_{1}^{2}(t)+\sigma_{1}^{2}(t)\right) \mu_{2}(t)  \tag{5.34}\\
E\left(X_{1}^{2}(t) X_{4}(t)\right)= & 2 \mu_{1}(t) \sigma_{14}(t)+\left(\mu_{1}^{2}(t)+\sigma_{1}^{2}(t)\right) \mu_{4}(t)  \tag{5.35}\\
E\left(X_{1}(t) X_{2}^{2}(t)\right)= & 2 \mu_{2}(t) \sigma_{21}(t)+\left(\mu_{2}^{2}(t)+\sigma_{2}^{2}(t)\right) \mu_{1}(t)  \tag{5.36}\\
E\left(X_{1}(t) X_{2}(t) X_{4}(t)\right)= & \mu_{1}(t) \mu_{2}(t) \mu_{4}(t)+\sigma_{24}(t) \mu_{1}(t) \\
& +\sigma_{14}(t) \mu_{2}(t)+\sigma_{12}(t) \mu_{4}(t) . \tag{5.37}
\end{align*}
$$

The expressions for $\boldsymbol{\mu}(t) \boldsymbol{G}^{T}(t)$ and $\boldsymbol{G}(t) \boldsymbol{G}^{T}(t)$ are

$$
\begin{gather*}
\boldsymbol{\mu}(t) \boldsymbol{G}^{T}(t)=\left(\begin{array}{ll}
\mu_{1}(t) E\left(X_{1}(t) X_{2}(t)\right) & \mu_{1}(t) E\left(X_{1}(t) X_{4}(t)\right) \\
\mu_{2}(t) E\left(X_{1}(t) X_{2}(t)\right) & \mu_{2}(t) E\left(X_{1}(t) X_{4}(t)\right) \\
\mu_{4}(t) E\left(X_{1}(t) X_{2}(t)\right) & \mu_{4}(t) E\left(X_{1}(t) X_{4}(t)\right)
\end{array}\right)  \tag{5.38}\\
\boldsymbol{G}(t) \boldsymbol{G}^{T}(t)=\left(\begin{array}{cc}
\left(E\left(X_{1}(t) X_{2}(t)\right)\right)^{2} & E\left(X_{1}(t) X_{2}(t)\right) E\left(X_{1}(t) X_{4}(t)\right) \\
E\left(X_{1}(t) X_{2}(t)\right) E\left(X_{1}(t) X_{4}(t)\right) & \left(E\left(X_{1}(t) X_{4}(t)\right)\right)^{2}
\end{array}\right) . \tag{5.39}
\end{gather*}
$$

The simulation results from the RM method were compared with the SSA method for the following parameters: $X_{1}(t=0)=10,000, X_{2}(t=0)=4,000, X_{4}(t=0)=1,000$, $c_{1}=0.0001 \mathrm{sec}^{-1}, c_{2}=0.0002 \mathrm{sec}^{-1}$ and $t=1.0 \mathrm{sec}$. In the RM method, $\Delta t=0.001 \mathrm{sec}$. The noisy results from the SSA method in Figure 5.1 are due to insufficient realizations of 1,000 . With 10,000 realizations in the SSA method (Figure 5.2), the results are less noisy and the RM method is in very good agreement with the SSA method.

The CPU time with the RM method on the Fermat machine is 0.09 sec and with the SSA method it is 1529 sec for 10,000 realizations. For a value of $\epsilon=0.001$, we obtain $\Delta t=0.001 \mathrm{sec}$ at time $t=0$. As the CPU time with the RM method is very small, we do not vary the time step during the course of the simulation.

### 5.5.2 Example - Dimerising Reaction

In the following reaction

$$
\begin{equation*}
R_{1}: 2 S_{1} \longrightarrow S_{2}, \tag{5.40}
\end{equation*}
$$

the monomer $S_{1}$ dimerises to $S_{2}$. We track species $S_{1}$, and therefore $N=1$, the number of reactions $M=1$ and the state of the system is given by $X_{1}(t)$. We also have

$$
\begin{equation*}
\nu_{11}=-2, \quad w_{11}=-2 c_{1} \Delta t, \quad h_{1}(t)=\frac{1}{2} X_{1}(t)\left(X_{1}(t)-1\right) \tag{5.41}
\end{equation*}
$$

and

$$
\begin{equation*}
r_{11}(t)=2 X_{1}(t)\left(X_{1}(t)-1\right) p_{11}\left(1-p_{11}\right) \tag{5.42}
\end{equation*}
$$

where $p_{11}=1-c_{1} \Delta t$. The expressions for the first and the second moments are

$$
\begin{align*}
\mu_{1}(t+\Delta t)= & \mu_{1}(t)+c_{1} \Delta t \mu_{1}(t)-c_{1} \Delta t E_{\boldsymbol{X}(t)}\left(X_{1}^{2}(t)\right)  \tag{5.43}\\
E_{\boldsymbol{X}(t+\Delta t)}\left(X_{1}^{2}(t+\Delta t)\right)= & E\left(X_{1}^{2}(t)\right)+2 E_{\boldsymbol{X}(t)}\left(X_{1}^{2}(t)\right) c_{1} \Delta t\left(1-c_{1} \Delta t\right) \\
& -2 \mu_{1}(t) c_{1} \Delta t\left(1-c_{1} \Delta t\right)-2 c_{1} \Delta t E\left(X_{1}^{2}(t)\left(X_{1}(t)-1\right)\right) \\
& +\left(c_{1} \Delta t\right)^{2} E\left(X_{1}^{2}(t)\left(X_{1}(t)-1\right)^{2}\right) \tag{5.44}
\end{align*}
$$

The simulation of this reaction and its comparison with other methods will be depicted in Chapter 6.

### 5.5.3 Example - Ten Species

We now illustrate an example with the number of species $N=10$, the number of reactions $M=8$. The reactions and the rate constants are listed in Table 5.2 and the initial molecular populations are given in Table 5.3. Results are displayed at time $t=0.5 \mathrm{sec}$ in Figure 5.3 and at time $t=2.0 \mathrm{sec}$ in Figure 5.4. In the RM method, $\Delta t=0.01 \mathrm{sec}$, and in the SSA method 2, 000 realizations were used. The results from the two methods are in good agreement.

### 5.5.4 Example - Viral Kinetics

We simulate the Viral Kinetics example described in Section 1.3.1 [73]. The species are $m R N A, D N A, \operatorname{Pr}$ (protein) and $V r$ (progeny virus). The reactions and the rate constants are given in Table 5.4. For the reaction constants given in this table, the molecular populations have the steady state values 20,200 and 10000 for the species $m R N A, D N A$ and $\operatorname{Pr}$, respectively. We start simulations with molecular populations of 2,200 and 10,000 for $m R N A, D N A$ and $\operatorname{Pr}$, respectively. It takes several days for the mRNA population to reach its steady state value. The time step in the RM method is 0.0001 day. The simulations results are displayed in Figures 5.5 and 5.6, at time $t=1.0$ day, using 80, 000 and 150, 000 realizations, respectively, in the SSA method. The CPU time on the Fermat machine is 0.8 sec using the RM method. Due to very large variance, the plots from SSA simulations are very noisy. The simulation times using the SSA method is about 16 hours with 80,000 realizations and 31 hours with 150,000 realizations. Figure 5.7 displays the results using the RM method at time $t=20.0$ days.

| Reaction | Reaction Rate <br> $\left(\mathrm{sec}^{-1}\right)$ |
| :--- | :--- |
| $S_{1}+S_{2} \longrightarrow S_{3}$ | 0.001 |
| $S_{1}+S_{4} \longrightarrow S_{5}$ | 0.002 |
| $S_{3} \longrightarrow S_{1}+S_{2}$ | 0.01 |
| $S_{5} \longrightarrow S_{1}+S_{4}$ | 0.01 |
| $S_{5} \longrightarrow S_{6}$ | 0.0001 |
| $S_{6} \longrightarrow S_{7}+S_{8}$ | 0.0003 |
| $S_{7} \longrightarrow S_{9}$ | 0.05 |
| $S_{8} \longrightarrow S_{10}$ | 0.05 |

Table 5.2: The reactions and rate constants for the Ten Species example in Section 5.5.3.

| Species | Initial Population |
| :--- | :--- |
| $S_{1}$ | 1000 |
| $S_{2}$ | 400 |
| $S_{3}$ | 0 |
| $S_{4}$ | 200 |
| $S_{5}$ | 0 |
| $S_{6}$ | 1000 |
| $S_{7}$ | 400 |
| $S_{8}$ | 500 |
| $S_{9}$ | 0 |
| $S_{10}$ | 0 |

Table 5.3: Initial molecular populations for the Ten Species example in Section 5.5.3.

| Reaction | ReactionRate <br> $\left(\right.$ day $\left.^{-1}\right)$ |
| :--- | :--- |
| $D N A \longrightarrow m R N A$ | 1.0 |
| $m R N A \longrightarrow D N A+m R N A$ | 0.025 |
| $m R N A \longrightarrow P r+m R N A$ | 1000 |
| $m R N A \longrightarrow \varnothing$ | 0.25 |
| $P r \longrightarrow \varnothing$ | 1.9985 |
| $D N A+P r \longrightarrow V r$ | $7.5 \times 10^{-6}$ |

Table 5.4: The reactions and rate constants for the Viral Kinetics example in Section 5.5.4.


Figure 5.1: Simulation results from the RM (solid line) and the SSA ("noisy" plot) methods for the Competing Reaction system in Section 5.5.1. The 1, 000 realizations used in the SSA method are not adequate.


Figure 5.2: Same example as in Figure 5.1, except that the SSA method utilized 10, 000 realizations, and therefore the plots are less "noisy".


Figure 5.3: Simulation results from the RM (solid line) and the SSA ("noisy" plot) methods, at time $t=0.5 \mathrm{sec}$, for the Ten Species example described in Section 5.5.3.


Figure 5.4: Simulation results from the RM (solid line) and the SSA ("noisy" plot) methods, at time $t=2.0 \mathrm{sec}$, for the Ten Species example described in Section 5.5.3.


Figure 5.5: Simulation results from the RM (solid line) and the SSA ("noisy" plot) methods, at time $t=1$ day, for the Viral Kinetics example described in Section 5.5.4. The number of realizations in the SSA method is 80,000 .


Figure 5.6: Same example as in Figure 5.5, except that the SSA method utilized 150, 000 realizations.


Figure 5.7: Viral Kinetics simulation, using the RM method, at time $t=20$ days.

## Chapter 6

## Simulation and Comparison with other methods

### 6.1 Introduction

As described in Chapter 3, the most popular methods of obtaining the time evolution of the joint probability density in a biochemical system are the Monte Carlo methods, such as the stochastic simulation algorithm and the $\tau$-leap methods.

The moment propagation methods, in which we trace a few lower order moments of $P(\mathrm{x}, t)$, were briefly introduced by Érdi [15] and van Kampen [76]. In 1992, Gillespie utilized the propagator of the Markov process to derive the moment evoluation equations. However, he describes his approximation procedure for moment closure as "lengthy and tedious" (Appendix C, [24]). The procedure consists of solving several sets of differential equations, truncated at different orders, and comparing successive solutions. In their 2007 paper, Gómez-Uribe et al. derived equations to track the means, variances and the covariances, starting from the CME [32]. They term their equations Mass Fluctuation Kinetics (MFK), and they present several examples to illustrate their method. In their 2008 paper, the authors separate the slow and fast reactions to obtain reduced slow and fast CMEs [33]. However, in evaluating the expressions for the second moments, they ignore the third moments and therefore their results are not as accurate as the SSA (Figures 2 and 3, [32]). In 2005, Hespanha et al. [38] and Singh et al. [67] presented a Stochastic Hybrid Systems (SHS)
method to compute the first few moments of $P(\mathrm{x}, t)$, and a moment closure technique to evaluate the higher order moments that arise in the expressions for these moments. They illustrate several examples with a few species and reactions. Several other approaches have been presented, such as the Linear Noise Approximation [43] and utilizing the moment generating function [21]. Since the SHS method is better developed than the other methods mentioned above, we will compare the Proposed Method with the SHS.

### 6.2 Stochastic Hybrid Systems Method

A hybrid system is a dynamical system consisting of discrete and continuous states, where the continuous state is described by a differential equation and the discrete states reset the continuous state according to some pre-specified conditions [37], [49]. A stochastic hybrid system (SHS) is a hybrid system where both the continuous and the discrete states maybe random [19], [39]. For example, the continuous state maybe described by a stochastic differential equation and the transitions between the continuous states are triggered by a Poisson process. We now present the stochastic hybrid system described by Hespanha et al. and reproduce their results for the evolution of a few lower-order moments. Their work is based on piecewise deterministic Markov process (PDMP) introduced by Davis [10]. A PDMP consists of a continuous state described by a deterministic differential equation, and the transition between continuous states described by a stochastic jump process, with state dependent transition probabilities. Davis gives a complete characterization of the extended generator of such processes, which is relevant for computing the necessary expectations.

For the biochemical system defined in Section 4.2 with $N$ molecular species and $M$ reactions, the SHS consists of the following:

- A hybrid state $(\boldsymbol{X}(t), \boldsymbol{\theta}(t))$, where $\boldsymbol{X}(t):[0, \infty) \rightarrow \mathbb{R}^{N}$ is a stochastic process called the continuous state, $\boldsymbol{\theta}(t):[0, \infty) \rightarrow \boldsymbol{\Theta}$ is a jump process called the discrete state, $\mathbb{R}^{N}$ denotes the $N$-dimensional Euclidean space and $\boldsymbol{\Theta}$ denotes a discrete set.
- A differential equation which governs $\boldsymbol{X}(t)$

$$
\begin{equation*}
\frac{d \boldsymbol{X}(t)}{d t}=g(\boldsymbol{\theta}, \boldsymbol{X}, t), \quad g: \boldsymbol{\Theta} \times \mathbb{R}^{N} \times[0, \infty) \rightarrow \mathbb{R}^{N} \tag{6.1}
\end{equation*}
$$

- A family of reset maps

$$
\begin{equation*}
\boldsymbol{\phi}_{m}\left(\boldsymbol{\theta}^{-}, \boldsymbol{X}^{-}, t\right), \quad \boldsymbol{\phi}_{m}: \boldsymbol{\Theta} \times \mathbb{R}^{N} \times[0, \infty) \rightarrow \boldsymbol{\Theta} \times \mathbb{R}^{N} \tag{6.2}
\end{equation*}
$$

for the $m$ th reaction defined by the stoichiometry of the species for that reaction. These maps determine how the continuous states are reset.

- A family of transition intensities

$$
\begin{equation*}
\lambda_{m}(\boldsymbol{\theta}, \boldsymbol{X}, t)=c_{m} h_{m}(\boldsymbol{X}, t), \quad \lambda_{m}: \boldsymbol{\Theta} \times \mathbb{R}^{N} \times[0, \infty) \rightarrow[0, \infty) \tag{6.3}
\end{equation*}
$$

for the $m$ th reaction, where $c_{m}$ is the stochastic rate constant and $h_{m}(\boldsymbol{X}, t)$ is the number of distinct molecular combinations of reaction $R_{m}$ as defined in Section 1.4.2.

For the above biochemical system, the continuous state evolves as follows:

$$
\begin{equation*}
\frac{d \boldsymbol{X}(t)}{d t}=0 \tag{6.4}
\end{equation*}
$$

This is because the molecular populations change only when there is a reaction, and these are set by the reset map defined in (6.2).

In a polynomial stochastic hybrid system ( $p S H S$ ), the reset maps $\boldsymbol{\phi}_{m}(\boldsymbol{\theta}, \boldsymbol{X}, t)$ and the transition intensities $\lambda_{m}(\boldsymbol{\theta}, \boldsymbol{X}, t)$ are all polynomial functions of $\boldsymbol{X}(t)$. Let the joint moments about the origin be given by

$$
\begin{equation*}
\gamma^{\boldsymbol{q}}(t) \equiv \gamma^{\left(q_{1}, q_{2} \ldots q_{N}\right)}(t)=E\left(X_{1}^{q_{1}}(t) X_{2}^{q_{2}}(t) \ldots X_{N}^{q_{N}}(t)\right), \tag{6.5}
\end{equation*}
$$

where $\boldsymbol{q}=\left(q_{1} q_{2} \ldots q_{N}\right) \in \mathbb{N}^{N}$ and $\mathbb{N}$ is the set of positive integers. Recall that the order of the moments $Q$ is defined by $Q=\sum_{n=1}^{N} q_{n}$. The authors show that the time evolution of $\gamma^{\boldsymbol{q}}(t)$ is described by

$$
\begin{equation*}
\frac{d \gamma^{\boldsymbol{q}}(t)}{d t}=E\left(\sum_{m=1}^{M}\left(\boldsymbol{\phi}_{m}^{\boldsymbol{q}}(\boldsymbol{\theta}, \boldsymbol{X}, t)-\boldsymbol{X}^{\boldsymbol{q}}(t)\right) \lambda_{m}(\boldsymbol{\theta}, \boldsymbol{X}, t)\right) \tag{6.6}
\end{equation*}
$$

In the above equation, $\boldsymbol{X}^{q}(t)$ denotes the monomial $X_{1}^{q_{1}}(t) X_{2}^{q_{2}}(t) \ldots X_{N}^{q_{N}}(t)$.

### 6.2.1 Moment Closure

If the biochemical system consists of second order reactions, then the expressions for moments of order up to $Q$ consists of moments of order greater than $Q$. The authors approximate the higher order moments as nonlinear functions of the lower order moments using a moment closure technique [67], [68]. The results for computing these nonlinear functions are presented here. For detailed derivations, we refer you to their papers.

Let $\gamma(t)=\left(\begin{array}{llll}\gamma^{\boldsymbol{q}_{1}}(t) & \gamma^{\boldsymbol{q}_{2}}(t) & \ldots & \gamma^{\boldsymbol{q}_{J}}(t)\end{array}\right)^{T}$ be a vector which contains all moments of order up to $Q$ and $\bar{\gamma}(t)=\left(\gamma^{\bar{q}_{1}}(t) \gamma^{\overline{\boldsymbol{q}}_{2}}(t) \ldots \gamma^{\overline{\boldsymbol{q}}_{J^{\prime}}}(t)\right)^{T}$ be a vector which contains moments of order greater than $Q$, required for the evaluation of the moments in $\gamma(t)$. The sizes of $\gamma(t)$ and $\bar{\gamma}(t)$ are $J$ and $J^{\prime}$, respectively. The time evolution of $\gamma$, omitting the $t^{\prime} s$, maybe expressed as

$$
\begin{equation*}
\frac{d \boldsymbol{\gamma}}{d t}=\boldsymbol{V}_{1} \boldsymbol{\gamma}+\boldsymbol{V}_{2} \overline{\boldsymbol{\gamma}} \tag{6.7}
\end{equation*}
$$

where the matrices $\boldsymbol{V}_{1}$ and $\boldsymbol{V}_{2}$ are determined from (6.6). The above system is closed by approximating the elements of $\bar{\gamma}$ as nonlinear functions of the elements of $\gamma$. Let $\boldsymbol{v}$ denote the new approximate system. Then the truncated moment dynamics is given by

$$
\begin{equation*}
\frac{d \boldsymbol{v}}{d t}=\boldsymbol{V}_{1} \boldsymbol{v}+\boldsymbol{V}_{2} \boldsymbol{\varphi}(\boldsymbol{v}) \tag{6.8}
\end{equation*}
$$

where $\boldsymbol{\varphi}(\boldsymbol{v})$ is the vector of moment closure function. Its elements, denoted by $\varphi(\boldsymbol{v})$, are determined by matching the time derivatives of $\boldsymbol{\gamma}$ and $\boldsymbol{v}$ with deterministic initial conditions. That is

$$
\begin{equation*}
\boldsymbol{\gamma}\left(t_{0}\right)=\left.\boldsymbol{v}\left(t_{0}\right) \Longrightarrow \frac{d^{k} \boldsymbol{\gamma}}{d t^{k}}\right|_{t=t_{0}}=\left.\frac{d^{k} \boldsymbol{v}}{d t^{k}}\right|_{t=t_{0}}, \quad \forall k=1,2, \ldots K \tag{6.9}
\end{equation*}
$$

where $K$ is selected to be large enough such that $\gamma$ remains "close" to $\boldsymbol{v}$. The authors assume that $\varphi(\boldsymbol{v})$ have the separable form

$$
\begin{equation*}
\varphi(\boldsymbol{v}) \equiv \varphi(\boldsymbol{\gamma})=\left(\gamma^{\boldsymbol{q}_{1}}\right)^{\eta_{1}}\left(\gamma^{\boldsymbol{q}_{2}}\right)^{\eta_{2}} \ldots\left(\gamma^{\boldsymbol{q}_{J}}\right)^{\eta_{J}} \tag{6.10}
\end{equation*}
$$

where $\gamma^{\boldsymbol{q}_{j}}, j=1 \ldots J$, are the elements of $\boldsymbol{\gamma}$ and $\eta_{1}, \eta_{2}, \ldots \eta_{J}$ are determined as described below. For any two vectors $\boldsymbol{q}=\left(q_{1}, q_{2} \ldots q_{N}\right)$ and $\overline{\boldsymbol{q}}=\left(\bar{q}_{1}, \bar{q}_{2} \ldots \bar{q}_{N}\right)$, define

$$
\begin{equation*}
\binom{\boldsymbol{q}}{\overline{\boldsymbol{q}}}=\binom{q_{1}}{\overline{q_{1}}}\binom{q_{2}}{\overline{q_{2}}} \ldots\binom{q_{N}}{\overline{q_{N}}} . \tag{6.11}
\end{equation*}
$$

Then for each element $\gamma^{\overline{\boldsymbol{q}}_{j^{\prime}}}$ in $\bar{\gamma}, \eta_{1}, \eta_{2}, \ldots \eta_{J}$ are determined from the equations

$$
\begin{equation*}
\binom{\overline{\boldsymbol{q}}_{j^{\prime}}}{\boldsymbol{q}_{j}}=\sum_{n=1}^{J} \eta_{n}\binom{\boldsymbol{q}_{n}}{\boldsymbol{q}_{j}} \quad \forall j=1,2, \ldots J . \tag{6.12}
\end{equation*}
$$

The authors indicate that their moment closure technique, which gives expressions for higher order moments as functions of lower order moments are consistent with lognormal distributions (p. 5004 in [69]).

In their earlier papers (Eq. (14) in [38]), the authors maintain that in order to satisfy the conditions in (6.10) and (6.12), for every moment $\gamma^{\boldsymbol{q}_{j^{\prime}}}$ in $\overline{\boldsymbol{\gamma}}$, the polynomial $\sum_{i=1}^{\infty} u_{j^{\prime}, i} \boldsymbol{x}^{\boldsymbol{q}_{i}}$ must belong to the linear subspace generated by the polynomials

$$
\begin{equation*}
\sum_{i=1}^{\infty} u_{j, i} \boldsymbol{x}^{\boldsymbol{q}_{j^{\prime}}-\boldsymbol{q}_{j}+\boldsymbol{q}_{i}}, \quad j=1 \ldots J \tag{6.13}
\end{equation*}
$$

However, in their recent paper, the authors have dropped the condition in (6.13) [69]. We illustrate the method with an example.

### 6.2.2 Example - Dimerising Reaction

Consider the dimerising reaction

$$
\begin{equation*}
R_{1}: 2 S_{1} \longrightarrow S_{2} \tag{6.14}
\end{equation*}
$$

The state of the system is $\boldsymbol{X}(t)=X_{1}(t)$. Omitting the $t$ 's, the reset map for the reaction is $\phi_{1}(\boldsymbol{X})=X_{1}-2$ and the transition intensity is $\lambda_{1}(\boldsymbol{X})=\frac{c_{1}}{2} X_{1}\left(X_{1}-1\right)$.

Let $\boldsymbol{\gamma}=\left(\begin{array}{ll}\gamma^{1} & \gamma^{2}\end{array}\right)^{T}$ and $\bar{\gamma}=\left(\gamma^{3}\right)$. From (6.6), we obtain the expressions for the time evolution of $\gamma$

$$
\binom{\frac{d \gamma^{1}}{d t}}{\frac{d \gamma^{2}}{d t}}=\left(\begin{array}{cc}
c_{1} & -c_{1}  \tag{6.15}\\
-2 c_{1} & 4 c_{1}
\end{array}\right)\binom{\gamma^{1}}{\gamma^{2}}+\binom{0}{-2 c_{1}}\left(\gamma^{3}\right) .
$$

In addition, we have, for the time evolution of $\gamma^{3}$

$$
\begin{equation*}
\frac{d \gamma^{3}}{d t}=4 c_{1} \gamma^{1}-10 c_{1} \gamma^{2}+9 c_{1} \gamma^{3}-3 c_{1} \gamma^{4} \tag{6.16}
\end{equation*}
$$

The expressions in (6.15) and (6.16) do not satisfy the condition in (6.13) as the polynomial

$$
\begin{equation*}
p_{3}(x)=4 c_{1} x-10 c_{1} x^{2}+9 c_{1} x^{3}-3 c_{1} x^{4} \tag{6.17}
\end{equation*}
$$

does not belong to the subspace generated by the polynomials

$$
\begin{align*}
& p_{1}(x)=c_{1} x^{3}-c_{1} x^{4}  \tag{6.18}\\
& p_{2}(x)=-2 c_{1} x^{2}+4 c_{1} x^{3}-2 c_{1} x^{4} \tag{6.19}
\end{align*}
$$

The expressions in (6.15) and (6.16) are modified by dropping the lower order moments

$$
\begin{gather*}
\binom{\frac{d \gamma^{1}}{d t}}{\frac{d \gamma^{2}}{d t}}=\left(\begin{array}{cc}
c_{1} & -c_{1} \\
0 & 4 c_{1}
\end{array}\right)\binom{\gamma^{1}}{\gamma^{2}}+\binom{0}{-2 c_{1}}\left(\gamma^{3}\right) .  \tag{6.20}\\
\frac{d \gamma^{3}}{d t}=9 c_{1} \gamma^{3}-3 c_{1} \gamma^{4} . \tag{6.21}
\end{gather*}
$$

The condition in (6.13) is now satisfied.
Further, an expression for $\gamma^{3}$ is obtained using the moment closure technique in (6.10) and (6.12)

$$
\begin{equation*}
\gamma^{3}=\left(\gamma^{1}\right)^{\eta_{1}}\left(\gamma^{2}\right)^{\eta_{2}} \tag{6.22}
\end{equation*}
$$

where $\eta_{1}$ and $\eta_{2}$ are obtained by solving the equations

$$
\begin{align*}
\eta_{1}+2 \eta_{2} & =3  \tag{6.23}\\
\eta_{2} & =3 \tag{6.24}
\end{align*}
$$

This yields $\eta_{1}=-3$ and

$$
\begin{equation*}
\gamma^{3}=\left(\frac{\gamma^{2}}{\gamma^{1}}\right)^{3} \tag{6.25}
\end{equation*}
$$

### 6.3 Comparison of the RM Method with the SHS Method

We now compare the expressions for the evolution of the first two moments from the RM method with the SHS method. We also compare the simulation results from the two methods and discuss implementation issues.

### 6.3.1 First Order Reactions

For the first order reactions, the evolution of moments is closed, that is, the computation of moments of order up to $Q$ do not require any higher order moments. We compare the expressions for moments obtained from the RM method with those obtained from the SHS method.

## Example - Three Species (continued)

Continuing the example presented in Section 4.4.1, we have the reactions

$$
\begin{align*}
& R_{1}: S_{1} \xrightarrow{c_{1}} S_{2}  \tag{6.26a}\\
& R_{2}: S_{2} \xrightarrow{c_{2}} S_{3} . \tag{6.26b}
\end{align*}
$$

The state of the system is $\boldsymbol{X}(t)=\left(X_{1}(t) \quad X_{2}(t) \quad X_{3}(t)\right)^{T}$. Omitting the $t$ 's, the reset maps for the two reactions are

$$
\boldsymbol{\phi}_{1}(\boldsymbol{X})=\left(\begin{array}{c}
X_{1}-1  \tag{6.27}\\
X_{2}+1 \\
X_{3}
\end{array}\right), \quad \phi_{2}(\boldsymbol{X})=\left(\begin{array}{c}
X_{1} \\
X_{2}-1 \\
X_{3}+1
\end{array}\right)
$$

The transition intensities are

$$
\begin{equation*}
\lambda_{1}(\boldsymbol{X})=c_{1} X_{1}, \quad \lambda_{2}(\boldsymbol{X})=c_{2} X_{2} . \tag{6.28}
\end{equation*}
$$

From (6.6), we have the following expression for evaluating moments

$$
\begin{align*}
\frac{d \gamma^{q}}{d t}= & E\left(c_{1} X_{1}\left(\left(X_{1}-1\right)^{q_{1}}\left(X_{2}+1\right)^{q_{2}} X_{3}^{q_{3}}-X_{1}^{q_{1}} X_{2}^{q_{2}} X_{3}^{q_{3}}\right)+\right. \\
& \left.c_{2} X_{2}\left(X_{1}^{q_{1}}\left(X_{2}-1\right)^{q_{2}}\left(X_{3}+1\right)^{q_{3}}-X_{1}^{q_{1}} X_{2}^{q_{2}} X_{3}^{q_{3}}\right)\right) \tag{6.29}
\end{align*}
$$

We track a total of nine first and second order moments, and therefore $J=9$. The required expressions from the SHS method are

$$
\begin{align*}
& \frac{d \gamma^{(1,0,0)}}{d t}=-c_{1} \gamma^{(1,0,0)}  \tag{6.30}\\
& \frac{d \gamma^{(0,1,0)}}{d t}=c_{1} \gamma^{(1,0,0)}-c_{2} \gamma^{(0,1,0)}  \tag{6.31}\\
& \frac{d \gamma^{(0,0,1)}}{d t}=c_{2} \gamma^{(0,1,0)}  \tag{6.32}\\
& \frac{d \gamma^{(2,0,0)}}{d t}=c_{1} \gamma^{(1,0,0)}-2 c_{1} \gamma^{(2,0,0)}  \tag{6.33}\\
& \frac{d \gamma^{(0,2,0)}}{d t}=c_{1} \gamma^{(1,0,0)}+c_{2} \gamma^{(0,1,0)}+2 c_{1} \gamma^{(1,1,0)}-2 c_{2} \gamma^{(0,2,0)}  \tag{6.34}\\
& \frac{d \gamma^{(0,0,2)}}{d t}=c_{2} \gamma^{(0,1,0)}+2 c_{2} \gamma^{(0,1,1)}  \tag{6.35}\\
& \frac{d \gamma^{(1,1,0)}}{d t}=-c_{1} \gamma^{(1,0,0)}+c_{1} \gamma^{(2,0,0)}-\left(c_{1}+c_{2}\right) \gamma^{(1,1,0)}  \tag{6.36}\\
& \frac{d \gamma^{(1,0,1)}}{d t}=-c_{1} \gamma^{(1,0,1)}+c_{2} \gamma^{(1,1,0)}  \tag{6.37}\\
& \frac{d \gamma^{(0,1,1)}}{d t}=c_{1} \gamma^{(1,0,1)}+c_{2} \gamma^{(0,2,0)}-c_{2} \gamma^{(0,1,1)}-c_{2} \gamma^{(0,1,0)} \tag{6.38}
\end{align*}
$$

Applying (4.17) and (4.39), the RM method yields the following expressions for first moments

$$
\begin{align*}
& \mu_{1}(t+\Delta t)=e^{-c_{1} \Delta t} \mu_{1}(t)  \tag{6.39}\\
& \mu_{2}(t+\Delta t)=\left(1-e^{-c_{1} \Delta t}\right) \mu_{1}(t)+e^{-c_{2} \Delta t} \mu_{2}(t)  \tag{6.40}\\
& \mu_{3}(t+\Delta t)=\left(1-e^{-c_{2} \Delta t}\right) \mu_{2}(t)+\mu_{3}(t) \tag{6.41}
\end{align*}
$$

And from (4.27), the RM method yields the following expressions for the second moments

$$
\begin{align*}
& E\left(X_{1}^{2}(t+\Delta t)\right)=E\left(X_{1}(t)\right) e^{-c_{1} \Delta t}\left(1-e^{-c_{1} \Delta t}\right)+E\left(X_{1}^{2}(t)\right)\left(e^{-c_{1} \Delta t}\right)^{2}  \tag{6.42}\\
& E\left(X_{1}(t+\Delta t) X_{2}(t+\Delta t)\right)=-E\left(X_{1}(t)\right) e^{-c_{1} \Delta t}\left(1-e^{-c_{1} \Delta t}\right) \\
&+E\left(X_{1}^{2}(t)\right) e^{-c_{1} \Delta t}\left(1-e^{-c_{1} \Delta t}\right) \\
&+E\left(X_{1}(t) X_{2}(t)\right) e^{-c_{1} \Delta t} e^{-c_{2} \Delta t} \tag{6.43}
\end{align*}
$$

We rewrite (6.33) from the SHS method, replacing the differentials with finite differences

$$
\begin{equation*}
E\left(X_{1}^{2}(t+\Delta t)\right)=E\left(X_{1}^{2}(t)\right)+c_{1} \Delta t E\left(X_{1}(t)\right)-2 c_{1} \Delta t E\left(X_{1}^{2}(t)\right) \tag{6.44}
\end{equation*}
$$

We also rewrite (6.42) from the RM method, using the Taylor series expansion of $e^{-c_{1} \Delta t}$

$$
\begin{align*}
E\left(X_{1}^{2}(t+\Delta t)\right) \approx & E\left(X_{1}(t)\right)\left(c_{1} \Delta t-\frac{3}{2}\left(c_{1} \Delta t\right)^{2}+\ldots\right)+ \\
& E\left(X_{1}^{2}(t)\left(1-2 c_{1} \Delta t+2\left(c_{1} \Delta t\right)^{2}\right)+\ldots\right) . \tag{6.45}
\end{align*}
$$

The above expression contains additional second and higher order terms in $c_{1} \Delta t$, compared with (6.44), which entails shorter time steps and consequently longer computation times with the SHS method.

We compare the simulation results obtained by the SHS method with the RM and the SSA methods, for the same parameters as in Section 4.4.1. The differential equations in the SHS method were solved using the Euler method. We recapitulate that the initial number of molecules are $\boldsymbol{X}(t=0)=\left(\begin{array}{ll}10000500 & 0\end{array}\right)^{T}$ and the stochastic rate constants are $c_{1}=0.1$ $\mathrm{sec}^{-1}$ and $c_{2}=1.0 \mathrm{sec}^{-1}$. The species are tracked from $t=0 \mathrm{sec}$ to $t=10 \mathrm{sec}$, with $\Delta t=0.005$ sec for the RM method, $\Delta t=0.00001 \mathrm{sec}$ for the SHS method and the number of realizations being 5,000 for the SSA method. The average number of time steps with the SSA method is 12,721 . From Figure 6.1, we see that the estimated moments obtained by the RM and SHS methods are in good agreement. In Figure 6.2, the time step in the SHS method was changed to $\Delta t=0.0005$, whereas the parameters of the other methods remained unchanged. With the larger time step, the SHS method does not perform as well.

### 6.3.2 Second Order Reactions

For the second order reactions, the computation of moments of order up to $Q$ require moments of order greater than $Q$. Therefore, in addition to determining the moment evolution equations from (6.6), we also require expressions for the higher order moments, which are evaluated from (6.10) and (6.12). We compare the dimerising reaction example.

### 6.3.3 Example - Dimerising Reaction (continued)

In the SHS method, we will use expression (6.15) instead of (6.20), as the authors dropped condition (6.13) in a recent paper [69]. The expressions for the first and the second moments are

$$
\begin{align*}
\frac{d \gamma^{1}}{d t} & =c_{1} \gamma^{1}-c_{1} \gamma^{2}  \tag{6.46}\\
\frac{d \gamma^{2}}{d t} & =-2 c_{1} \gamma^{1}+4 c_{1} \gamma^{2}-2 c_{1} \gamma^{3} \tag{6.47}
\end{align*}
$$

Rewriting the above expressions, replacing the differentials with finite differences, yields

$$
\begin{align*}
\mu_{1}(t+\Delta t)= & \mu_{1}(t)+c_{1} \Delta t \mu_{1}(t)-c_{1} \Delta t E\left(X_{1}^{2}(t)\right)  \tag{6.48}\\
E\left(X_{1}^{2}(t+\Delta t)\right)= & E\left(X_{1}^{2}(t)\right)-2 c_{1} \Delta t \mu_{1}(t) \\
& +4 c_{1} \Delta t E\left(X_{1}^{2}(t)\right)-2 c_{1} \Delta t E\left(X_{1}^{3}(t)\right) \tag{6.49}
\end{align*}
$$

where the expression for $E\left(X_{1}^{3}(t)\right)$ is given by (6.25).
In the RM method, we rewrite expression (5.43) for the mean and expand expression (5.44) for the second moment

$$
\begin{equation*}
\mu_{1}(t+\Delta t)=\mu_{1}(t)+c_{1} \Delta t \mu_{1}(t)-c_{1} \Delta t E\left(X_{1}^{2}(t)\right) \tag{6.50}
\end{equation*}
$$

$$
\begin{align*}
E\left(X_{1}^{2}(t+\Delta t)\right)= & E\left(X_{1}^{2}(t)\right)-2 c_{1} \Delta t \mu_{1}(t) \\
& +4 c_{1} \Delta t E\left(X_{1}^{2}(t)\right)-2 c_{1} \Delta t E\left(X_{1}^{3}(t)\right) \\
& +2\left(c_{1} \Delta t\right)^{2} \mu_{1}(t)-\left(c_{1} \Delta t\right)^{2} E\left(X_{1}^{2}(t)\right) \\
& -\left(c_{1} \Delta t\right)^{2} E\left(X_{1}^{3}(t)\right)+\left(c_{1} \Delta t\right)^{2} E\left(X_{1}^{4}(t)\right), \tag{6.51}
\end{align*}
$$

where $E\left(X_{1}^{3}(t)\right)$ and $E\left(X_{1}^{4}(t)\right)$ are evaluated from

$$
\begin{align*}
& E\left(X_{1}^{3}(t)\right)=\mu_{1}^{3}(t)+3 \mu_{1}(t) \sigma_{1}^{2}(t)  \tag{6.52}\\
& E\left(X_{1}^{4}(t)\right)=\mu_{1}^{4}(t)+6 \mu_{1}^{2}(t) \sigma_{1}^{2}(t)+3 \sigma_{1}^{4}(t) \tag{6.53}
\end{align*}
$$

Once again, the expression for $E\left(X_{1}^{2}(t+\Delta t)\right)$ from the RM method contains additional terms, compared with the SHS method.

Figure 6.3 compares the simulation results, from the two methods, with the SSA. The parameters are $X_{1}(t=0)=1,500, X_{2}(t=0)=0, c_{1}=0.002 \mathrm{sec}^{-1}$ and the simulation is performed up to time $t=1 \mathrm{sec}$. The number of realizations in the SSA method is 3,000 . In the RM method $\Delta t=0.002 \mathrm{sec}$. The results from the SHS method compare well with the SSA with $\Delta t=0.0001 \mathrm{sec}$ (top plot), but not with the larger $\Delta t=0.002 \mathrm{sec}$ (bottom plot). This is due to the additional higher order terms in the RM method, and different moment closure expressions in the two methods. The RM method utilizes the normal distribution for the moment closure approximation, whereas the moment closure expressions in the SHS method correspond to lognormal distribution.

### 6.3.4 Implementation Issues

We now discuss the implementation issues with the SHS method, for examples consisting of a large number of molecules and species, such as the first order complex example presented in 4.4.2 and the second order example with ten species presented in 5.5.3.

In order to execute the complex first order example given in Section 4.4.2, with the number of species $N=100$ and the number of reactions $M=200$, using the SHS method, the expression (6.6) for the evolution of moments will consist of 200 terms. Many of these terms would be zero. Nonetheless, we would require a total of $N$ equations for the first moments and $N(N+1) / 2$ equations for the second moments. This gives a total of 5,150
equations similar to those in (6.30) through (6.38). This is much more arduous to implement than the RM method. For a given biochemical system with first order reactions, the RM method employs expressions (4.6a) and (4.6b) for computing $\boldsymbol{P}$ and (4.25a) and (4.25b) for computing $\boldsymbol{B}(t)$. The mean and the covariance at each time step are obtained using the recursive expressions in (4.17) and (4.29).

In order to execute the second order example given in Section 5.5.3, with the number of species $N=10$ and the number of reactions $M=8$, using the SHS method, we have to derive a total of 55 equations, for the evolution of the first and the second moments. In addition we have to solve the system of equations in (6.12), to determine the expressions for higher order moments. These are many more additional steps compared to the RM method. For a given system of second order reactions, the RM method is implemented as described in Section 5.4. That is, given the reaction network and rate constants, define matrices $\boldsymbol{W}$, $\boldsymbol{X}(t)$ and $\boldsymbol{H}(t)$, compute the higher order expectations utilizing Table 5.1, and compute $\boldsymbol{\mu}(t+\Delta t)$ and $\boldsymbol{C}(t+\Delta t)$, recursively, by employing expressions (5.8) and (5.14), without deriving or solving any differential equations.

To recapitulate, the RM method produces the same accuracy as the SHS method, in fewer time steps, due to additional higher order terms and a different moment closure technique. The moment closure techniques from various methods require further investigation. For a given biochemical system, with known reactions and rate constants, the RM method can be implemented by utilizing the given recursive expressions, whereas the SHS method requires additional steps or additional software to derive the differential equations, determine the expressions for higher order moments and solve the resulting equations.


Figure 6.1: Comparison of the results from the RM, SHS and the SSA methods for the Three Species example given in Section 6.3.1. The curves represent the computed means $\mu_{2}(t)$ (top plot), the variances $\sigma_{2}^{2}(t)$ (second plot), and the covariances $\sigma_{21}(t)$ (third plot) and $\sigma_{23}(t)$ (bottom plot). $\Delta t$ is 0.005 sec with the RM method and 0.00001 sec with the SHS method.


Figure 6.2: Same example and parameters as in Figure 6.1, except that in the SHS method $\Delta t=0.0005 \mathrm{sec}$. With a larger $\Delta t$, the SHS method does not agree well with the RM and the SSA methods.


Figure 6.3: Simulation results from the RM (solid line), the SHS (dotted line) and the SSA ("noisy" plot) methods for the Dimerising Reaction example described in Section 6.3.3. In the top plot, $\Delta t$ is 0.002 sec in the RM method and 0.0001 sec in the SHS method. When $\Delta t$ is changed to 0.002 sec in the SHS method (bottom plot), it doesn't agree very well with the SSA.

## Chapter 7

## Conclusion and Future Work

In the forward problem of biochemical reaction systems, we have knowledge of all the reactions and the associated rate constants. We want to determine the joint probability distribution of the populations of all the molecular species at any time instant. The present state-of-the-art approaches for stochastic simulation of such systems are based on Monte Carlo methods, such as the Stochastic Simulation Algorithm (SSA) and its accelerated versions. The Monte Carlo methods provide approximations of the complete distribution, but they are computationally infeasible for complex biochemical systems consisting of a large number of species and reactions.

An alternate approach is to determine the evolution of a few lower order moments with time. In this dissertation, we presented a new method, called the Recursive Moment (RM) method, for propagating the first two moments of the joint probability distribution of the molecular populations in complex biochemical systems. The proposed method is applicable to very large systems. Compared with the SSA, the RM method yields significant savings in computation time. In contrast to other moment propagation methods, the recursive expressions in our method can be implemented by specifying rate constants and reaction stoichiometries, without having to derive or solve any differential equations.

An important direction for continued work on the RM method includes evaluation of its numerical accuracy. We would like to be in a position to predict the accuracy of the proposed method. Therefore, it is important to study the accuracy as a function of various factors, including the value of the time step, the values of the reaction constants, the types of reactions, and the population sizes of the species. Ideally, we would like to be able
to determine upper bounds of performance given a particular set of reactions, molecular populations, and choices of time step.

There are cases where the distributions of the molecular populations are not unimodal. More specifically, there may be scenarios where the PMFs may be multimodal, for example, the bimodal distribution in the Schlögl model [66]. This implies that tracking only the first two moments will not be sufficient. Thus, it is important to expand the work presented in the dissertation so that one can track higher order moments. The approach used for deriving the proposed method is amenable to generalize the method. The goal, however, is to obtain a procedure that is not cumbersome for execution. Clearly, once we have higher order moments, we will be able to deduce much more about the joint distribution of the species.

An important extension of the proposed work is to apply the method in practice. To that end, simulations of complex metabolic and signaling pathways, such as the cell signaling pathway in cancer and stem cells as described in [13], are particularly attractive. Since such pathways are models developed by biologists, it would be interesting to test them against experimental data. Therefore, developing ways of testing proposed models using the simulation tools would add considerable value to the analysis.

Finally, in the problem studied here, it was assumed that we have knowledge of the reaction network and all the rate constants. However, these quantities are usually unknown, but what is available instead are the time series data of biochemical pathways. Further, it is difficult to obtain data for all the species sampled at high enough frequencies, and one may only have observations of a subset of the species. In future work, one would like to examine statistical techniques for inferring parameters from data in challenging scenarios. The class of Bayesian methods offer approaches for tackling many difficult problems [6], [18]. With partially observed data and for systems with a large number of species (high dimension), computational techniques such as the Markov Chain Monte Carlo (MCMC) can be used to solve such problems [64].

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## Appendix A

## Acronyms

CME
FPE
ODE
PDE
PMF
PDF
MGF
MFK
RM
RRE
SDE
SHS
SSA
chemical master equation
Fokker-Planck equation
ordinary differential equation
partial differential equation
probability mass function (discrete random variable)
probability density function (continuous random variable)
moment generating function
mass fluctuation kinetics
recursive method
reaction rate equation
stochastic differential equation
stochastic hybrid system
stochastic simulation algorithm

## Appendix B

## Notation

| $a_{m}(\boldsymbol{x}, t)$ | reaction propensity in $R_{m}$. |
| :--- | :--- |
| $\alpha_{i}$ | population of species $S_{i}$ at time $t=0$. |
| $\boldsymbol{\alpha}$ | state of the system at time $t=0$. |
| $b(\cdot)$ | SDE drift coefficient. |
| $c_{m}$ | stochastic rate constant of reaction $R_{m}$. |
| $C_{i}$ | sum of the stochastic rate constants in which species $S_{i}$ |
|  | is a source. |
| $\boldsymbol{C}(t)$ | covariance of $\boldsymbol{X}(t)$ |
| $\delta(t)$ | Dirac delta function. |
| $\Delta t$ | step size of the recursive method. |
| $\mathcal{G}\left(s_{i}, t\right)$ | probability generating function of $P\left(x_{i}, t\right)$. |
| $\mathcal{G}(\boldsymbol{s}, t)$ | probability generating function of $P(\mathrm{x}, t)$. |
| $\gamma^{\boldsymbol{q}}(t)$ | elements of $\gamma(t)$ |
| $\gamma^{\prime \boldsymbol{q}}(t)$ | elements of $\gamma^{\prime}(t)$ |
| $\gamma(t)$ | vector of non-central moments |
| $\gamma^{\prime}(t)$ | vector of central moments |
| $\Gamma(\cdot)$ | Gamma function. |
| $h_{m}(\boldsymbol{x}, t)$ | number of distinct molecular combinations in $R_{m}$. |
| $k_{m}$ | deterministic rate constant of reaction $R_{m}$. |

$\mathcal{M}_{\boldsymbol{X}(t)}(\boldsymbol{\omega})$
$\boldsymbol{\mu}(t)$,
$n_{A}$
$N$
$\mathbb{N}$
$\mathcal{N}\left(\mu, \sigma^{2}\right)$
$\nu_{m i}$
$\boldsymbol{\nu}_{m}$
$P(\mathrm{x}, t), P(\mathrm{x}(t))$
$P_{l \mid l-1}(\cdot)$
$Q$
$R_{m}$
$\mathbb{R}$
$S_{i}, S_{j}$
$t$ or $t_{l}$
$u(\cdot)$
V
$\mathcal{W}(\cdot)$
$\boldsymbol{x}(t)$
$X_{i}(t)$
$\boldsymbol{X}(t)$
number of reactions.
moment generating function of $\boldsymbol{X}(t)$.
mean of $\boldsymbol{X}(t)$
Avogadro number
number of species or the number of tracked species.
set of positive integers.
normal distribution with mean $\mu$ and variance $\sigma^{2}$.
stoichiometry of species $S_{i}$ in reaction $R_{m}$.
stoichiometry vector in reaction $R_{m}$.
pdf or pmf of of $\boldsymbol{X}(t)$.
conditional pmf of a random vector at time $t_{l}$ given its
pmf at time $t_{l-1}$.
order of the moments in $\gamma^{\boldsymbol{q}}(t)$.
$m$ th reaction, $m=1 \ldots M$.
set of real numbers.
$i$ th and $j$ th species, $i=1 \ldots N, j=1 \ldots N$.
time index
SDE diffusion coefficient.
volume of biochemical reaction system
Wiener process.
a realization of $\boldsymbol{X}(t)$.
population of species $S_{i}$ at time $t$.
state of the system at time $t$, with elements $X_{i}(t)$.

